

Exact recovery in graphical models

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Philippe Rigollet

- Exact recovery in the Ising blockmodel
Q. Berthet, P. Rigollet, and P. S.

[arXiv:1612.03880v2](#)

To appear in the *Annals of Statistics*

Motivation

- Finding communities in populations, based on similar **behavior** and **influence**.
- One of the justifications for **stochastic blockmodels**
- What if we observe the **behavior**, not the graph?

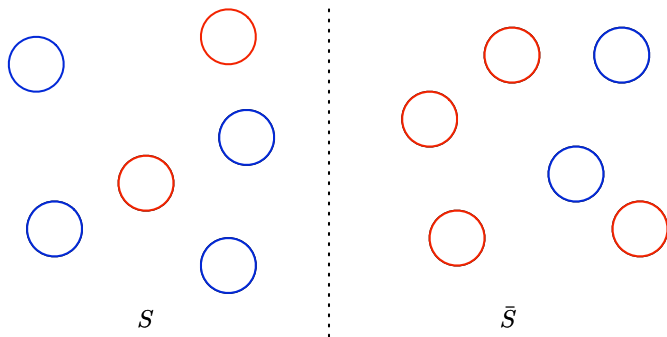
Motivation: Learning the model

- Population of interacting agents are often modeled as graphs
- **Setting:** The population has an (unknown) community structure
 - ▶ This is what the graph represents
- But one observes only **behavior** of the nodes, not the graph!

Behaviour \equiv "Sample from the 'model'"
Community structure \equiv "the underlying 'model'"

- Common framework in economics, biology (protein-protein interactions) ...
 - ▶ But one needs to be wary of application specific caveats

Motivation: the abstraction



Model with p individuals, $\sigma \in \{-1, 1\}^p$ and balanced communities (S, \bar{S}) .

$$\mathbf{P}_S(\sigma) = \underbrace{\hspace{10em}}_{?}.$$

Problem: Find S from observations of σ

Prior beliefs

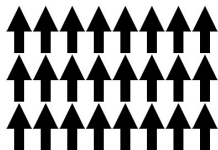
- To solve the problem, one needs to fix a search space of models
- Specifically, there needs to be some reasonably concrete model of the inter-community and intra-community interactions

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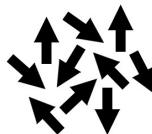
- To solve the problem, one needs to fix a search space of models
- Specifically, there needs to be some reasonably concrete model of the inter-community and intra-community interactions
- We choose to model them in a fashion inspired by model inspired from statistical mechanics

Magnets: the Ising model

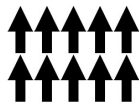
Some magnetic materials lose their magnetism just above a critical temperature (Curie temperature)



$$T < T_c$$



Applied Magnetic
Field Absent



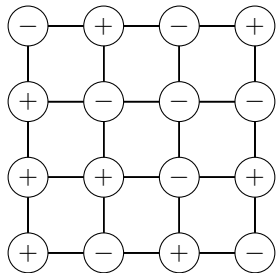
Applied Magnetic
Field Present

$$T > T_c$$

Perhaps the first phase transition to be abstracted: **Ising model** (1925)

Image credits: User ACGrain at English Wikipedia,

Ising model



Gibbs distribution

$$\mu(\sigma) \propto \exp \left(\beta \sum_{u \sim v} \sigma_u \sigma_v \right)$$

- Nodes “represent” magnetic domains, β represents inverse temperature $1/T$
- Mean magnetization: Average of $\sum_v \sigma_v$ according to μ

$$M(\beta) = E_{\sigma \sim \mu} \left[\sum_v \sigma_v \right]$$

Ising model: Beyond magnets

The Ising model[★] makes an appearance in a rather wide variety of areas:

- Individual choice theory, as the **logit response**

e.g. [McKelvey & Palfrey, *Games and Econ. Behaviour*, 1995]

- Spread of opinions in social networks

e.g. [Montanari & Saberi, *PNAS*, 2010]

- Computer vision

e.g. [Geman & Graffigne, *Proc. of the ICM*, 1986]

★ ...and other related models

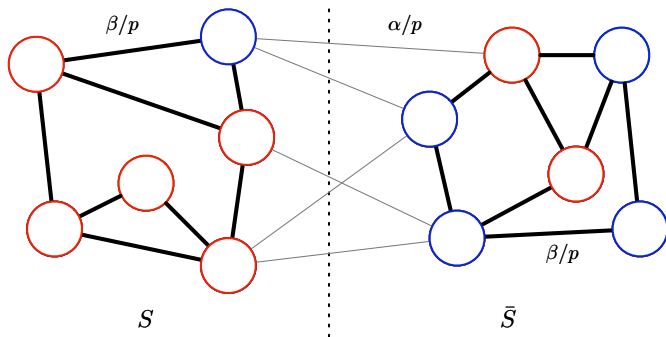
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- To solve the problem, one needs to fix a search space of models
- Specifically, there needs to be some reasonably concrete model of the inter-community and intra-community interactions
- We choose to model them in a fashion inspired by the **Ising model**

Recall the earlier cited applications to networks

e.g. [Montanari & Saberi, *PNAS*, 2010]

The Ising blockmodel



Model with p individuals, $\sigma \in \{-1, 1\}^p$ and balanced communities (S, \bar{S}) .

$$\mathbf{P}_S(\sigma) = \frac{1}{Z_{\alpha, \beta}} \exp \left[\frac{\beta}{2p} \sum_{i \sim j} \sigma_i \sigma_j + \frac{\alpha}{2p} \sum_{i \sim j} \sigma_i \sigma_j \right].$$

[Berthet, Rigollet, S., *Ann. Stat.*, to appear]

Problem description

Ising blockmodel:

$$\mathbf{P}_S(\sigma) = \frac{1}{Z_{\alpha,\beta}} \exp \left[\frac{\beta}{2p} \sum_{i \sim j} \sigma_i \sigma_j + \frac{\alpha}{2p} \sum_{i \not\sim j} \sigma_i \sigma_j \right], \quad \sigma \in \{-1, 1\}^p$$

- Balance: $|S| = |\bar{S}| = p/2$
- Homophily: $\beta > 0$
- Assortativity: $\beta > \alpha$

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Observations

$\sigma^{(1)}, \dots, \sigma^{(n)} \in \{-1, 1\}^p$ i.i.d. from \mathbf{P}_S

Objective: recover the *balanced* partition (S, \bar{S}) from observations

Problem overview

- Structure of the problem visible in the **covariance matrix** Σ

$$\Sigma = \mathbf{E}[\sigma\sigma^\top] = \left(\frac{\Delta}{\Omega} \middle| \frac{\Omega}{\Delta} \right) + (1 - \Delta)I_p.$$

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Main object of study

Scaling of $\Delta - \Omega$ with p ?

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Main object of study

Scaling of $\Delta - \Omega$ with p ?
Is it clear that $\Delta - \Omega$ is **positive**?

The role of the correlation matrix: Lower bound

- Consider patterns S and T that differ on exactly one pair of nodes
 - ▶ Let \mathbf{P}_S and \mathbf{P}_T be the induced distributions on samples

The role of the correlation matrix: Lower bound

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For an algorithm to be able to distinguish between \mathbf{P}_S and \mathbf{P}_T , they should be “far” in a statistical sense

The role of the correlation matrix: Lower bound

- Consider patterns S and T that differ on exactly one pair of nodes
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Lemma

$$\text{Relative entropy } D_{KL}(\mathbf{P}_T \| \mathbf{P}_S) = \frac{(\beta - \alpha)(p - 2)}{p} \cdot (\Delta - \Omega)$$



Theorem

$$\text{Sample complexity } n \gtrsim \frac{\log p}{\Delta - \Omega}$$

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- Note that the lemma implies that $\Delta - \Omega \geq 0$

The role of the covariance matrix: Upper bound

Empirical covariance matrix

$$\hat{\Sigma} = \frac{1}{n} \sum_{t=1}^n \sigma^{(t)} (\sigma^{(t)})^\top = \Sigma \pm O\left(\sqrt{\frac{\log p}{n}}\right) \quad \text{entrywise}$$

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Log likelihood

$$\mathcal{L}_{\sigma^{(1)}, \sigma^{(2)}, \dots, \sigma^{(n)}}(S) := \log \mathbf{P}_S(\sigma^{(1)}, \sigma^{(2)}, \dots, \sigma^{(n)}) = \text{const} + \frac{n(\beta - \alpha)}{2p} \mathbf{Tr} \left[\hat{\Sigma} v_S v_S^\top \right]$$

$$v_S := (\mathbb{I}_S - \mathbb{I}_{\bar{S}})$$

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Maximum likelihood estimator

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MINIMUM-BISECTION: NP-hard

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- Easier to work with projections on space \perp to $\mathbf{1}$:

$$\Gamma = P\Sigma P \quad \hat{\Gamma} = P\hat{\Sigma}P,$$

where

$$P = I - \frac{1}{p}\mathbf{1}\mathbf{1}^\top$$

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$$\hat{V} = v_S v_S^\top \iff L_S(\hat{\Gamma}) := \operatorname{diag}(\hat{\Gamma} v_S v_S^\top) - \hat{\Gamma} \succeq 0$$

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Spectrum of $L_S(\Gamma)$ ("infinite samples")

$$L_S(\Gamma) = \left(1 - \Delta + p \frac{\Delta - \Omega}{2}\right) \frac{\mathbf{1}_{[p]} \mathbf{1}_{[p]}^\top}{\sqrt{p} \sqrt{p}} + p \frac{\Delta - \Omega}{2} \cdot \tilde{I}_{\perp(\mathbf{1}, v_S)} + 0 \cdot v_S v_S^\top$$

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Thus, the relaxation is exact in the finite sample case if

$$\|L_S(\Gamma) - L_S(\hat{\Gamma})\|_{\text{op}} \leq p \frac{\Delta - \Omega}{2} \implies L_S(\hat{\Gamma}) \succeq 0 \implies \hat{V} = V_S$$

The role of the correlation matrix: Tight bounds

Upper and lower bounds

- The SDP relaxation is exact if $\|L_S(\Gamma) - L_S(\hat{\Gamma})\|_{\text{op}} \leq p(\Delta - \Sigma)/2$.
- From matrix concentration, this holds if

$$n \gtrsim \frac{\log p}{\Delta - \Omega}$$

- Previously, we saw $n \gtrsim \frac{\log p}{\Delta - \Omega}$ is required. Combining, we have:

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Theorem: Sample complexity and the correlation matrix

The sample complexity for exactly recovering the partition (S, \bar{S}) in the Ising blockmodel satisfies

$$n \simeq \frac{\log p}{\Delta - \Omega}$$

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Question

How does $\Delta - \Omega$ behave?

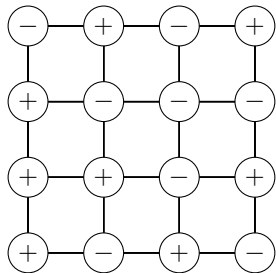
Analyzing $\Delta - \Omega$

Phase diagram of the Ising blockmodel

Ising blockmodel

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Ising model



Gibbs distribution

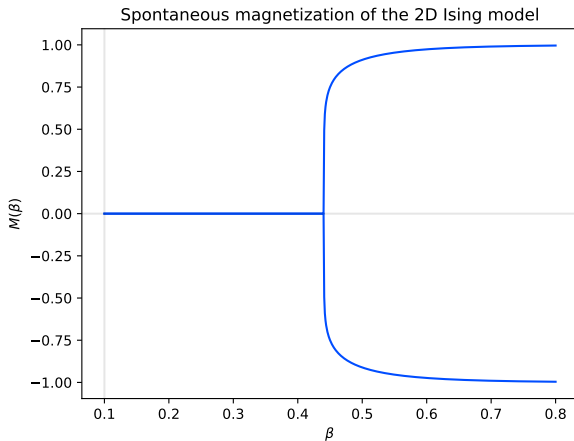
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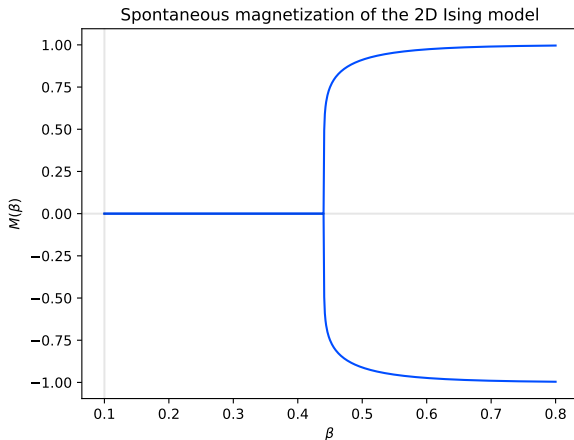
Phase transition in the Ising model

[Onsager, 1944]



Phase transition in the Ising model

[Onsager, 1944]



- Need to take an infinite volume limit to actually see this

Phase transition in the Ising model: finite volume

- But one can see a “phase transition” in some finite settings as well

Phase transition in the Ising model: finite volume

- But one can see a “phase transition” in some finite settings as well
- As the size of the “lattice” increases, the bimodal nature of μ becomes more pronounced

The Curie-Weiss model ($\alpha = \beta$)

"Mean field" Ising

- Mean magnetization: $\mu = \frac{\mathbf{1}^\top \sigma}{p} \in [-1, 1]$.

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- Free energy:

$$\mathbf{P}_\beta(\mu) \approx \frac{1}{Z_\beta} \exp \left(-\frac{p}{4} g_\beta^{\text{CW}}(\mu) \right), \quad g_\beta^{\text{CW}}(\mu) = -2\beta\mu^2 + 4h\left(\frac{1+\mu}{2}\right)$$

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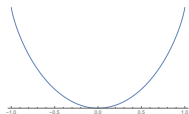
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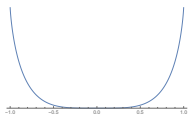
- Ground states: Minimizers $G \subset [-1, 1]$ of $g_\beta^{\text{CW}}(\mu)$.

Free energy of the Curie-Weiss model

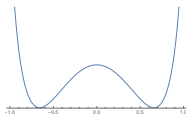
Ground states $\mathcal{G} = \{\tilde{\mu}(\beta), -\tilde{\mu}(\beta)\}, \tilde{\mu}(\beta) \geq 0$:



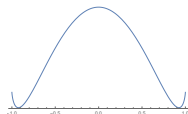
$\beta = 0$



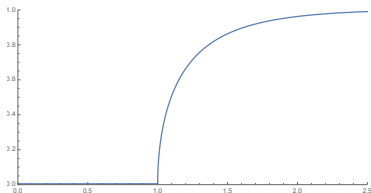
$\beta = 1$



$\beta = 1.2$



$\beta = 1.8$



$\beta \mapsto \tilde{\mu}(\beta)$

$$\Delta \approx \frac{1}{|G|} \sum_{s \in G} s^2 = \tilde{\mu}(\beta)^2$$

Free energy of the Ising blockmodel

- Energy is determined by mean magnetizations: $(\mu_S, \mu_{\bar{S}}) = \frac{2}{p}(\mathbf{1}_S^\top \sigma, \mathbf{1}_{\bar{S}}^\top \sigma)$

$$\mathbf{P}_S(\sigma) = \frac{1}{Z_{\alpha,\beta}} \exp\left(-\frac{p}{8}(-\beta\mu_S^2 - \beta\mu_{\bar{S}}^2 - 2\alpha\mu_S\mu_{\bar{S}})\right)$$

- **Marginal:** number of configurations with magnetizations μ is $\binom{(p/2)}{\frac{1+\mu}{2}(p/2)}$

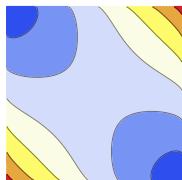
$$\mathbf{P}_S(\mu_S, \mu_{\bar{S}}) \approx \frac{1}{Z_{\alpha,\beta}} \exp\left(-\frac{p}{8} g_{\alpha,\beta}(\mu_S, \mu_{\bar{S}})\right)$$

where $g_{\alpha,\beta}$ is the free energy defined by

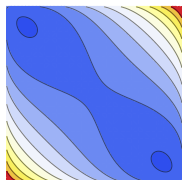
$$g_{\alpha,\beta}(\mu_S, \mu_{\bar{S}}) = -\beta\mu_S^2 - \beta\mu_{\bar{S}}^2 - 2\alpha\mu_S\mu_{\bar{S}} + 4h\left(\frac{1+\mu_S}{2}\right) + 4h\left(\frac{1+\mu_{\bar{S}}}{2}\right).$$

Ground states for the Ising blockmodel

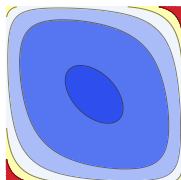
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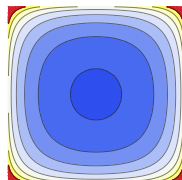
$$\alpha = -6$$



$$\alpha = -2.5$$

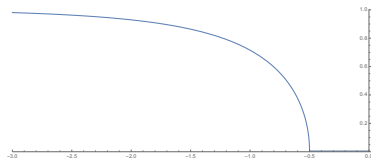


$$\alpha = -0.5$$



$$\alpha = 0$$

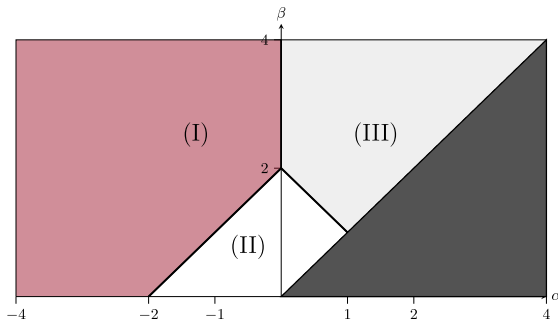
Ground states on the skew-diagonal ($\tilde{\mu}_S = -\tilde{\mu}_{\bar{S}}$) for $\alpha \leq 0$ and fixed $\beta = 1.5 < 2$



$$\alpha \mapsto \tilde{\mu}_S(\alpha, \beta = 1.5)$$

Phase diagram

Full understanding of the position of the ground states for $\beta > 0$, $\alpha < \beta$



Theorem

- Phase diagram for all the parameter regions
 - ▶ Region (I): Two ground states $(\tilde{\mu}_S, \tilde{\mu}_{\bar{S}}) = \pm(\tilde{x}, -\tilde{x})$
 - ▶ Region (II): One ground state at $(0, 0)$
 - ▶ Region (III): Two ground states $(\tilde{\mu}_S, \tilde{\mu}_{\bar{S}}) = \pm(\tilde{x}, \tilde{x})$

Concentration

- Quantities of interest as expectations

$$\Delta \approx \frac{1}{2} \mathbf{E}[\mu_S^2 + \mu_{\bar{S}}^2] \quad \text{and} \quad \Omega \approx \mathbf{E}[\mu_S \mu_{\bar{S}}].$$

Theorem: Gaussian approximation of the Gibbs distribution

For “nice” $\varphi : [-1, 1]^2 \rightarrow \mathbb{R}^+$,

$$\mathbf{E}_{\alpha, \beta}[\varphi(\mu)] \simeq_p \frac{1}{|G|} \sum_{\tilde{s} \in G} \mathbf{E}_{Z \sim \mathcal{N}(0, I)} \left[\varphi(\tilde{s} + 2\sqrt{\frac{2}{p}} H^{-1/2} Z) \right]$$

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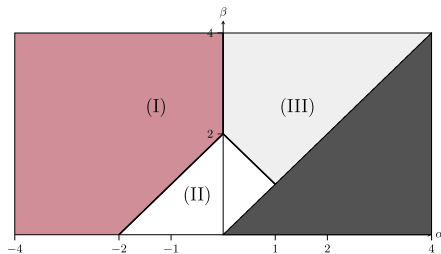
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- Applied to $\Delta - \Omega$:

$$\Delta - \Omega \simeq_p \begin{cases} 2\tilde{x}^2 (= \Theta(1)) & \text{in region (I)} \\ \frac{C_{\alpha, \beta}}{p} & \text{in region (II)} \\ \frac{C'_{\alpha, \beta}}{p} & \text{in region (III)} \end{cases}$$

Exact recovery



Recall that

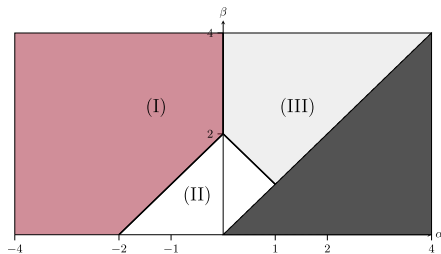
Theorem: Sample complexity and the correlation matrix

The sample complexity for exactly recovering the partition (S, \bar{S}) in the Ising blockmodel satisfies

$$n \simeq \frac{\log p}{\Delta - \Omega}$$

Combining the above estimates, we get...

Exact recovery



Main result: Optimal sample size for exact recovery

(I) : $n^* \approx \log(p)$ (II) and (III) : $n^* \approx \textcolor{red}{p} \log(p)$.

Conclusion

- Open questions
 - ▶ **Non-independent samples:** Samples drawn from a logit response/Glauber dynamics?
 - ▶ **Algorithmic results:** Can the SDP be circumvented?
 - ▶ Generalization to multiple blocks, more complex structures

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Stochastic blockmodels

- **one** observation of random graph on p vertices

$$\mathbf{P}(i \leftrightarrow j) = \begin{cases} b & \text{for all } i \sim j \\ a & \text{for all } i \not\sim j \end{cases}$$

- Exact recovery using SDP iff

$$a = \mathbf{a} \frac{\log p}{p}, b = \mathbf{b} \frac{\log p}{p}$$

and

$$(\mathbf{a} + \mathbf{b})/2 > 1 + \sqrt{\mathbf{a}\mathbf{b}}$$

Wigner matrices

Graphical models / MRF

- n observations $\sigma^{(1)}, \dots, \sigma^{(n)}$ i.i.d.

$$\mathbf{P}(\sigma) \propto \exp \left[\frac{\beta}{2p} \sum_{i,j} J_{ij} \sigma_i \sigma_j \right]$$

- Goal estimate sparse $J = \{J_{ij}\}_{ij}$ (max degree d)
- Sample complexity $n \gg 2^d \log p$

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Abbé, Bandeira, Hall '14

Hajek, Wu '16

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Chow-Liu '68

Bresler, Mossel, Sly '08

Santhanam, Wainwright '12

Bresler '15

Vuffray, Misra, Lokhov, Chertkov '16

Wishart matrices