

Experimentation with Temporal Interference: Poisson's Equation and Adaptive Markov Chain Sampling

Peter W. Glynn
Stanford University

Joint work with Ramesh Johari and Mohammad Rasouli

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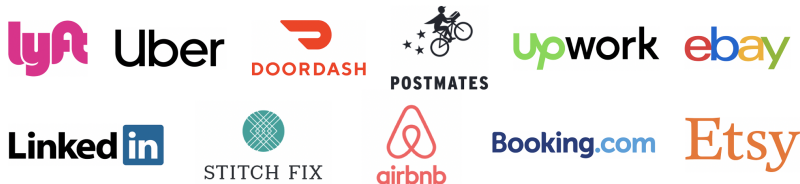
Outline of Talk

1. What is “temporal interference”?
2. Discussion of related concepts:
 - Poisson’s equation for Markov chains
 - Markov decision processes
3. Optimal adaptive Markov chain sampling

1. What is "Temporal Interference"?

Motivation: Testing Algorithms

Suppose you are one of these:



You have two algorithms A and B that you want to compare (e.g., matching algorithms).

Each algorithm changes the *state* of the system.

How do you design an experiment (A/B test) and an estimator to compare them?

Naive Solution: Randomize Over Time

Suppose at each decision epoch, we randomly flip a coin and run either A (heads) or B (tails).

Why is this not a good idea?

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Why is this not a good idea?

Temporal interference: Each algorithm's action changes the *state* as seen by the other algorithm.

Therefore experimental units (time steps) *interfere* with each other, introducing *bias*.

Industry Practice: Switchback Designs

Many platforms (ridesharing, delivery marketplaces, etc.) use *switchback designs* to run A/B tests of algorithms:

1. Divide time into *fixed length non-overlapping intervals*.
2. In each successive interval, assign one of algorithm A or B .
3. Compute sample average estimate $\widehat{\text{SAE}}_A$ and $\widehat{\text{SAE}}_B$ of reward of A and B respectively.
4. Compute $\widehat{\text{SAE}}_A - \widehat{\text{SAE}}_B$ as *treatment effect estimate* $\widehat{\text{TE}}$.



Note: Doesn't eliminate temporal interference.

Overview of Our Contributions

We cast the problem of testing two algorithms as a theoretical problem of *testing two Markov chains*.

We focus on *consistent* estimation of TE.

- We develop a *Markov policy* for allocation, that together with a MLE for \widehat{TE} , is *consistent* and *sample efficient*.
- We develop a *regenerative policy* for allocation that is *consistent* when used with the SAE for \widehat{TE} (but not sample efficient).

2. Discussion of Related Concepts:

- Poisson's Equation for Markov Chains
- Markov Decision Processes

Poisson's Equation for Markov Chains

- $X = (X_n : n \geq 0)$, S -valued Markov chain, irreducible, $|S| < \infty$
- $P = (P(x, y) : x, y \in S)$ transition matrix

Given a function/column vector f , Poisson's equation is

$$(P - I)g = -f$$

For solvability: Need $\pi f = 0$

Poisson's Equation for Markov Chains

$$(P - I)g = -f$$

$$(P - I + \Pi)g = -f$$

$$g = (I - P + \Pi)^{-1}f \quad (\triangleq (I - B)^{-1}f)$$

$$(I - P + \Pi)^{-1} = \sum_{n=0}^{\infty} (P - \Pi)^n \quad \text{aperiodic setting}$$

$$(I - P + \Pi)^{-1} \text{ exists in general}$$

Remark: $(I - P + \Pi)^{-1}$ is known as the *fundamental matrix*

Poisson's Equation for Markov Chains

An application:

- Suppose we want to prove

$$\frac{1}{n} \sum_{j=0}^{n-1} f(X_j) \xrightarrow{a.s.} Ef(X_\infty)$$

as $n \rightarrow \infty$

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- Try to write $\sum_{j=0}^{n-1} f(X_j) - nEf(X_\infty)$ in terms of a martingale
- Put $f_c(x) = f(x) - Ef(X_\infty)$ and solve

$$(P - I)g = -f_c$$

Poisson's Equation for Markov Chains

- Note that

$$E[g(X_i) \mid \mathcal{F}_{i-1}] = \sum_y P(X_{i-1}, y)g(y) = (Pg)(X_{i-1})$$

so

$$D_i = g(X_i) - (Pg)(X_{i-1})$$

is a martingale difference

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- But

$$\begin{aligned} M_n &= \sum_{i=1}^n [g(X_i) - (Pg)(X_{i-1})] \\ &= \sum_{i=0}^{n-1} [g(X_i) - (Pg)(X_i)] + g(X_n) - g(X_0) \\ &= \sum_{i=0}^{n-1} f_c(X_i) + g(X_n) - g(X_0) \quad (\text{recall: } (P - I)g = -f_c) \end{aligned}$$

Poisson's Equation for Markov Chains

- So,

$$\frac{1}{n} \sum_{i=0}^{n-1} f_c(X_i) = \frac{1}{n} M_n + \frac{1}{n} g(X_0) - \frac{1}{n} g(X_n)$$

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- Martingale theory:

$$\frac{1}{n} M_n \xrightarrow{a.s.} 0$$

Poisson's Equation for Markov Chains

The Central Limit Theorem (CLT) for Markov Chains:

$$\frac{1}{\sqrt{n}} \sum_{i=0}^{n-1} f_c(X_i) = \frac{1}{\sqrt{n}} M_n + \frac{1}{\sqrt{n}} g(X_0) - \frac{1}{\sqrt{n}} g(X_n)$$

- Martingale CLT implies:

$$\frac{1}{\sqrt{n}} \sum_{i=0}^{n-1} f_c(X_i) \Rightarrow \sigma N(0, 1)$$

where

$$\sigma^2 = \text{var}_\pi D_1 = E_\pi g^2(X_0) - E_\pi (Pg)^2(X_0)$$

Poisson's Equation for Markov Chains

Many other applications of Poisson's equation:

- stochastic control
- gradients of $E^\theta f(X_\infty)$
- non-stationary Markov chains

Markov Decision Processes

- Compute an optimal policy/control minimizing

$$\overline{\lim}_{n \rightarrow \infty} \frac{1}{n} \sum_{j=0}^{n-1} c(X_j, A_j)$$

over all adapted policies $(A_n : n \geq 0)$

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- The optimality equation is:

$$v(x) + \gamma = \min_a \left[c(x, a) + \sum_y P_a(x, y) v(y) \right]$$

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- Optimal policy: Choose action a in state x with probability

$$\frac{\pi(x,a)}{\sum_{a'} \pi(x,a')}$$

3. Optimal Adaptive Markov Chain Sampling

Non-parametric Model

- Discrete time $n = 0, 1, 2, \dots$
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At time n : State X_n , action A_n , reward R_n

The Estimation Problem

Treatment effect of interest is the *steady state reward difference*:

$$\begin{aligned}\alpha &= \alpha(2) - \alpha(1) = \sum_x \pi(2, x)r(2, x) - \sum_x \pi(1, x)r(1, x) \\ &= \boldsymbol{\pi}(2)\boldsymbol{r}(2) - \boldsymbol{\pi}(1)\boldsymbol{r}(1).\end{aligned}$$

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We get to choose an **estimator** and a **policy**:

- **Estimator**: $\alpha = (\alpha_n : n \geq 0), \alpha_n \in \mathbb{R}$
- **Policy**: $A = (A_n : n \geq 0), A_n \in \{1, 2\}$

Estimator and policy are adapted to history, and policy can be randomized.

The Non-parametric Maximum Likelihood Estimator

Definitions:

$$\Gamma_n(\ell, x) := \# \text{ of plays of } \ell \text{ in first } n \text{ steps} = \sum_{j=0}^{n-1} I(X_j = x, A_j = \ell)$$

$$r_n(\ell, x) := \text{SAE of } r(\ell, x) = \frac{\sum_{j=0}^{n-1} I(X_j = x, A_j = \ell) r(\ell, x)}{\max\{\Gamma_n(\ell, x), 1\}}$$

$$P_n(\ell, x, y) := \text{SAE of } P(\ell, x, y) = \frac{\sum_{j=0}^{n-1} I(X_j = x, A_j = \ell, X_{j+1} = y)}{\max\{\Gamma_n(\ell, x), 1\}}$$

Let $\pi_n(\ell)$ be invariant distribution of $P_n(\ell)$ (exists a.s. as $n \rightarrow \infty$).

Then:

$$\alpha_n^{\text{MLE}} = \pi_n(2)r_n(2) - \pi_n(1)r_n(1).$$

Time-Average Regular Policies

We optimize over time-average regular policies.

Definition

Policy A is *time-average regular* if

$$\frac{1}{n}\Gamma_n(\ell, x) \xrightarrow{p} \gamma(\ell, x)$$

as $n \rightarrow \infty$ for each $x \in S, \ell = 1, 2$, and (possibly random) $\gamma(\ell, x)$.

We call $\gamma = (\gamma(\ell, x) : x \in S, \ell = 1, 2)$ the *policy limit*.

(For our theory we will require $\gamma(\ell, x) > 0$ a.s.)

Central Limit Theorem for MLE

Theorem

For any time-average regular policy A with strictly positive policy limits:

$$n^{1/2}(\alpha_n^{\text{MLE}} - \alpha) \Rightarrow \sum_x \frac{\pi(2, x)\sigma(2, x)}{\gamma(2, x)^{1/2}} G(2, x) - \sum_x \frac{\pi(1, x)\sigma(1, x)}{\gamma(1, x)^{1/2}} G(1, x).$$

where:

- $G(\ell, x)$ are i.i.d. $N(0, 1)$;
- $\sigma^2(\ell, x) = \text{Var}(\tilde{g}(\ell, X_j) \mid X_{j-1} = x, A_{j-1} = \ell)$
(assume positive);
- $\tilde{g}(\ell)$ solves the following *Poisson equation*:

$$\tilde{g}(\ell) = (\mathbf{I} - \mathbf{P}(\ell) + \mathbf{\Pi}(\ell))^{-1} \mathbf{r}(\ell)$$

- $\mathbf{\Pi}(\ell)$ is the matrix where each row is $\pi(\ell)$.

Optimal Oracle Policy for MLE

Let \mathcal{K} be the (convex, compact) set of all $(\kappa(\ell, x) : x \in S, \ell = 1, 2)$ such that:

$$\kappa(1, y) + \kappa(2, y) = \sum_{\ell} \sum_x \kappa(\ell, x) P(\ell, x, y), \quad y \in S;$$

$$\sum_{\ell} \sum_x \kappa(\ell, x) = 1;$$

$$\kappa(\ell, x) \geq 0.$$

Lemma: The law of any time-average regular policy limit γ is a probability measure over \mathcal{K} .

Optimal Oracle Policy for MLE

Let κ^* be the solution to the following convex optimization problem:

$$\begin{aligned} & \text{minimize} && \sum_{\ell} \sum_x \frac{\pi^2(\ell, x) \sigma^2(\ell, x)}{\kappa(\ell, x)} \\ & \text{subject to} && \kappa \in \mathcal{K}. \end{aligned}$$

Then κ^* can be realized as the policy limit of the following *stationary, Markov* policy:

Run algorithm ℓ in state x with probability:

$$p^*(\ell, x) = \frac{\kappa^*(\ell, x)}{\kappa^*(1, x) + \kappa^*(2, x)}.$$

Optimal Oracle Policy for MLE

Theorem

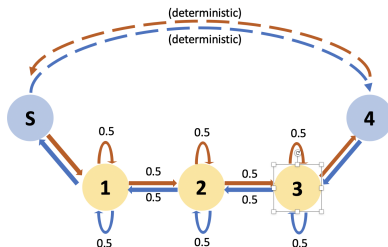
The policy p^* minimizes the asymptotic variance of $n^{1/2}(\alpha_n^{\text{MLE}} - \alpha)$ over time-average regular policies.

The Value of Cooperative Exploration

Cooperative exploration: Two chains can yield much more efficient estimation than either chain alone.

Example: Deterministic reward $r = 1$ in states 1, 2, 3, and zero reward elsewhere.

Estimating **red** or **blue** chain alone has asymptotic variance $\Theta(S)$ higher than using both together!



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where $g(i)$ is the solution to Poisson's equation for $r_c(i)$

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- $$((P_n(i) - P(i))g(i))(x) = \frac{\sum_{j=1}^n D_j(i, x)}{\pi_n(i, x)}$$

where

$$D_j(i, x) = I(X_{j-1} = x, A_{j-1} = i) [g(i, X_j) - (P(i)g(i))(X_{j-1})]$$

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- Now, apply martingale CLT

Optimal Online Policy for MLE

Without knowledge of the primitives, we can compute $\kappa_n(\ell, x)$ as the optimal solution given $\mathbf{P}_n(\ell)$, and set:

$$p_n(\ell, x) = (1 - \epsilon_n) \left(\frac{\kappa_n(\ell, x)}{\kappa_n(1, x) + \kappa_n(2, x)} \right) + \frac{1}{2}\epsilon_n,$$

with $\epsilon_n = n^{-1/2}$ (forced exploration).

This yields the asymptotically optimal policy limits in an online fashion.

Analysis easily extends to random rewards $R_n(i)$

Summary and Looking Ahead

We proposed a benchmark model with which to evaluate sampling efficiency of consistent estimator-design pairs for switchback experimentation.

There are several considerations we have not addressed:

- Finite horizon analysis
- Multiple treatments
- Nonstationarity

Three lectures hinting at the range of different problems of interest to the OR/MS applied probability community:

- output analysis for Monte Carlo
- Markov chains and processes in the presence of time-of-day effects, day-of-week effects, etc.
- A/B testing and temporal interference

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Just the “tip of the iceberg”

Thank you!