Bayesian model selection and parameter estimation

Bayesian inference: summary

• Parameter estimation: compute posterior density function $p(\vec{\theta}|d, H, I) \propto p(d|\vec{\theta}, H, I) p(\vec{\theta}|H, I)$

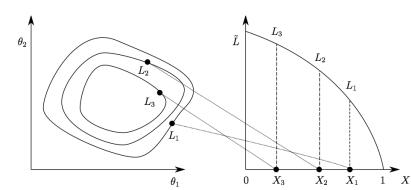
Posterior density function for one parameter, e.g. θ_1 :

$$p(heta_1|d,H,I) = \int_{ heta_2^{\min}}^{ heta_2^{\max}} \dots \int_{ heta_N^{\min}}^{ heta_N^{\max}} p(heta_1,\dots, heta_N|d,H,I) \, d heta_2\dots d heta_N$$

Model selection: compute odds ratio

$$\begin{split} O_{H_2}^{H_1} &\equiv \frac{p(H_1|d,I)}{p(H_2|d,I)} = \frac{p(d|H_1,I)}{p(d|H_2,I)} \frac{p(H_1|I)}{p(H_2|I)} \\ p(d|H,I) &= \int p(d|\vec{\theta},H,I) \, p(\vec{\theta}|H,I) \, d\vec{\theta} \end{split}$$

• Compute using e.g. nested sampling $p(d|H,I) = Z = \int \tilde{L}(X) dX \approx \sum_{k} L_k \Delta X_k$ $p(\vec{\theta}|d,H,I) = \tilde{P}(X) = \frac{\tilde{L}(X)}{Z} \Delta X$



Parameter space is 15-dimensional:

$$ec{ heta} = \{m_1, m_2, ec{S_1}, ec{S_2}, lpha, \delta, \iota, \psi, d_{ ext{L}}, t_c, arphi_c\}$$

Different detectors D have different response to signals:

$$ilde{h}^{(D)}(f) = \left[F^{(D)}_{+} ilde{h}_{+}(f) + F^{(D)}_{\times} ilde{h}_{\times}(f)\right] e^{-2\pi i f \Delta t^{(D)}}$$

where $F^{(D)}_{+}(\alpha, \delta, \psi, t_0)$ and $F^{(D)}_{\times}(\alpha, \delta, \psi, t_0)$ antenna pattern functions at geocentric arrival time t_0 while $\Delta t^{(D)}(\alpha, \delta, t_0)$ differences between arrival times at geocenter and at detectors

Different noise realizations in different detectors:

$$\tilde{d}^{(D)}(f) = \tilde{h}^{(D)}(f) + \tilde{n}^{(D)}(f)$$

Different noise power spectral densities:

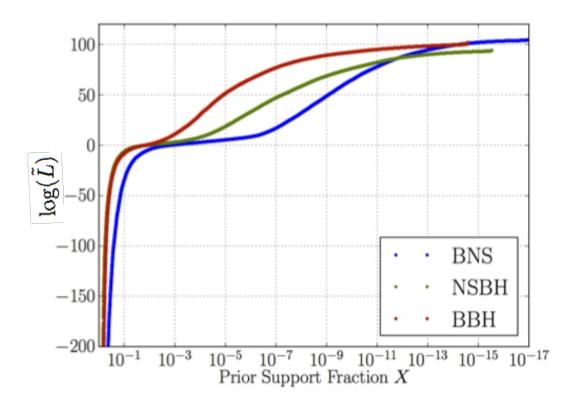
$$\langle \tilde{n}^{(D)}(f) \, \tilde{n}^{(D')^*}(f')
angle = rac{1}{2} \delta(f - f') \delta_{DD'} S^{(D)}(f)$$

- Examples of generative hypotheses:
 - Data is pure noise, \mathcal{H}_N
 - Data contains signal (with waveform from a certain family), \mathcal{H}_S
- Probability for noise realization n^(D) = n₀ is given by p(n^(D) = n₀) ∝ e^{-(n₀|n₀)/2}
 Likelihood for H_N: p(d^(D)|θ, H_N, I) ∝ e^{-(d^(D)|d^(D))/2}
- Likelihood for \mathcal{H}_S : $p(d^{(D)}|\vec{\theta}, \mathcal{H}_S, I) \propto e^{-(d^{(D)} - h^{(D)}|d^{(D)} - h^{(D)})/2}$ (Recall $\vec{d}^{(D)}(f) = \vec{h}^{(D)}(f) + \tilde{n}^{(D)}(f)$)
- Data from different detectors $\vec{d} = \{d^{(H)}, d^{(L)},\}$
- Joint likelihood: $p(\vec{d}|\vec{\theta}, \mathcal{H}, I) = \prod_{(D)} p(\vec{d}^{(D)}|\vec{\theta}, \mathcal{H}, I)$

- Assume that the search pipelines have found a compact binary coalescence signal
- Compute evidence for signal hypothesis \mathcal{H}_S :

$$p(d|\mathcal{H}_S, I) = Z = \int \tilde{L}(X) dX \approx \sum_k L_k \Delta X_k$$

• Typical growth of $\tilde{L}(X)$: usually convenient to consider logarithm



Posterior densities for parameters:

$$p(\vec{\theta}|d, \mathcal{H}_S, I) = \tilde{P}(X) = \frac{\tilde{L}(X)}{Z} \Delta X$$

Marginalize to get posterior density for one particular parameter:

$$p(heta_1|d,\mathcal{H}_S,I) = \int_{ heta_2^{\min}}^{ heta_2^{\max}} \dots \int_{ heta_N^{\min}}^{ heta_N^{\max}} p(heta_1,\dots, heta_N|d,\mathcal{H}_S,I)$$

... or get posterior density for two parameters jointly:

$$p(heta_1, heta_2|d,\mathcal{H}_S,I) = \int_{ heta_3^{\min}}^{ heta_3^{\max}} \dots \int_{ heta_N^{\min}}^{ heta_N^{\max}} p(heta_1,\dots, heta_N|d,\mathcal{H}_S,I)$$

Often useful to change to a different set of parameters:

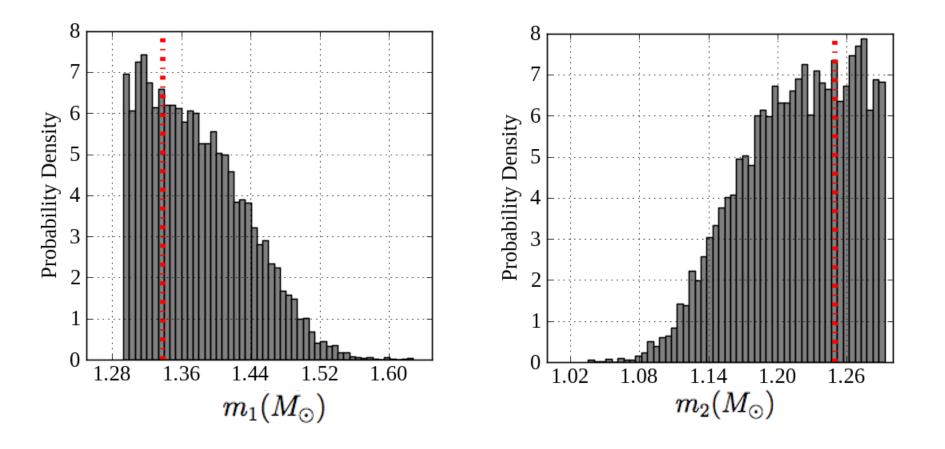
$$\begin{split} p(\vec{\theta}|d,\mathcal{H}_S,I) \, d\vec{\theta} \;&=\; p(\vec{\lambda}|d,\mathcal{H}_S,I) \, d\vec{\lambda} \\ \;&=\; p(\vec{\lambda}|d,\mathcal{H}_S,I) \det \frac{\partial(\lambda_1,\ldots,\lambda_N)}{\partial(\theta_1,\ldots,\theta_N)} \, d\vec{\theta} \end{split}$$

hence

$$p(ec{ heta}|d,\mathcal{H}_S,I)=p(ec{\lambda}|d,\mathcal{H}_S,I)\detrac{\partial(\lambda_1,\ldots,\lambda_N)}{\partial(heta_1,\ldots, heta_N)}$$

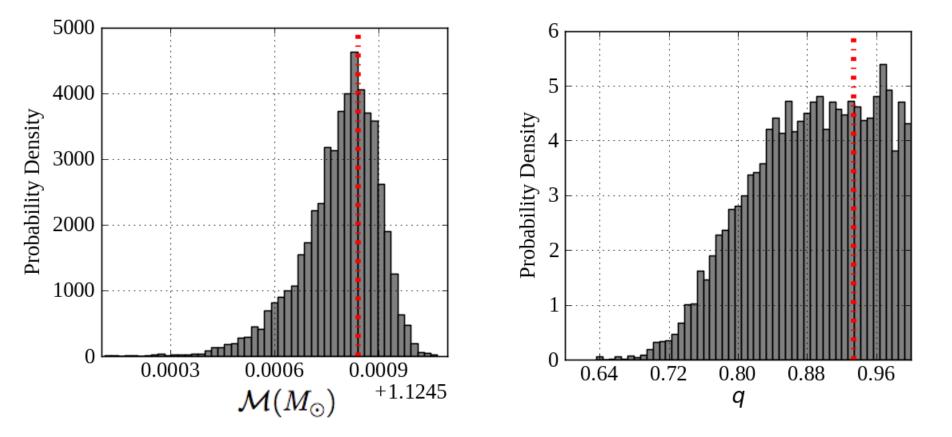
Example: binary neutron star coalescence, $(m_1, m_2) = (1.34, 1.25)M_{sun}$ at distance of 156.5 Mpc, simulated Gaussian noise for Advanced LIGO/Virgo

Component masses:



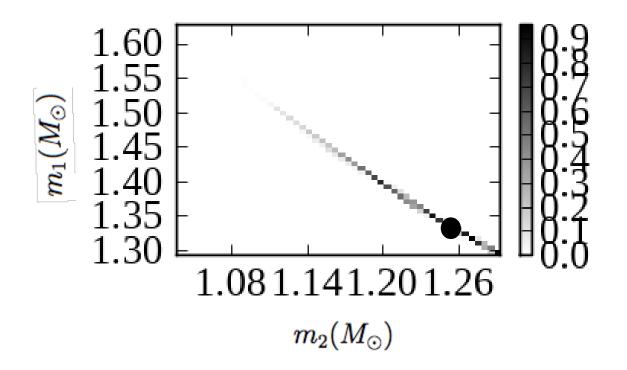
Example: binary neutron star coalescence, $(m_1, m_2) = (1.34, 1.25)M_{sun}$ at distance of 156.5 Mpc, simulated Gaussian noise for Advanced LIGO/Virgo

• Chirp mass
$$\mathcal{M}=(m_1+m_2)^{-1/5}(m_1m_2)^{3/5}$$
 and mass ratio $q=m_2/m_1$:



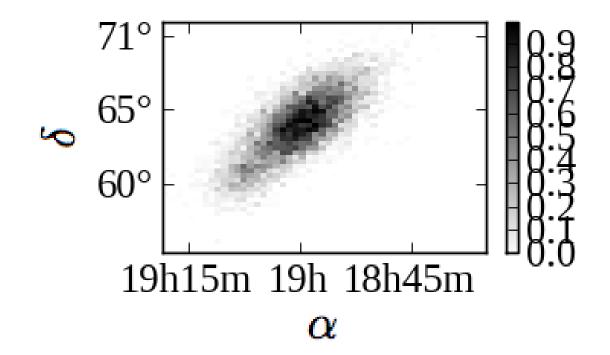
Example: binary neutron star coalescence, $(m_1, m_2) = (1.34, 1.25)M_{sun}$ at distance of 156.5 Mpc, simulated Gaussian noise for Advanced LIGO/Virgo

• Joint posterior density distribution for (m_1, m_2) :



Example: binary neutron star coalescence, $(m_1, m_2) = (1.34, 1.25)M_{sun}$ at distance of 156.5 Mpc, simulated Gaussian noise for Advanced LIGO/Virgo

Position on the sky:



<u>Combining information from multiple sources</u>

- Masses, sky position, distance, ... are incidental (though it will be of great interest to see how they are distributed!)
- Will sometimes want to check a *functional dependence* E.g.:
 - Dependence of post-Newtonian parameters on masses, spins
 - Neutron star equation of state $\text{P}(\rho)$
- Two ways of doing this:
 - Hypothesis testing: compare different possible functional dependences
 - Measure parameters that determine functional dependence
- Assuming the functional dependence is universal, should be able to combine information from multiple sources

Example: the equation of state of neutron stars

- As binary neutron stars spiral towards each other, they start feeling each other's tidal effects
- Quadrupole deformation induced in one star by tidal tensor of the other:
 $Q_{ij} = -\lambda(m) \, \mathcal{E}_{ij}$

Tidal deformability $\lambda(m)$ depends on neutron star equation of state $P(\rho)$

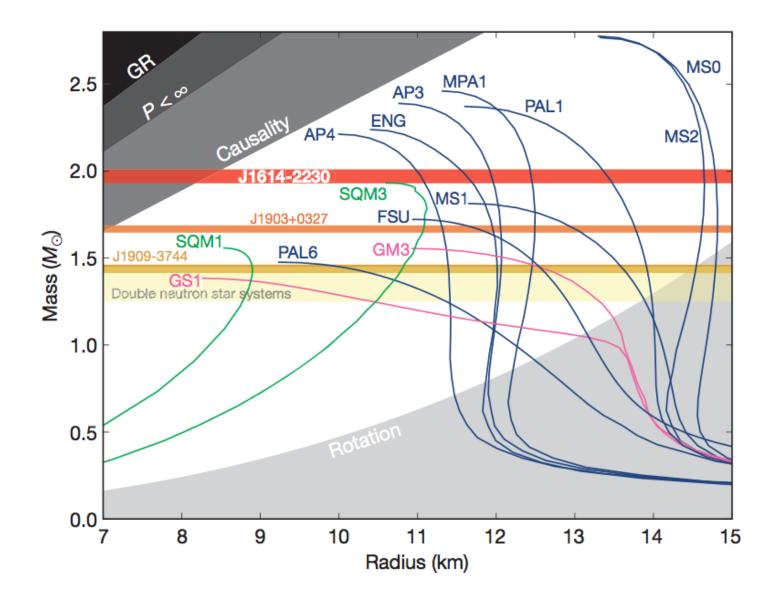
- Neutron star deformations affect surrounding spacetime curvature \rightarrow Effect on orbital motion, e.g. angular motion $\Phi(t)$ \rightarrow Imprinted onto the phase of the gravitational wave signal $2\Phi(t)$
- Contributions to the phase don't show up until 5PN order in phase but with large prefactor:

 $\lambda(m)/M^5 \propto (R/M)^5 \sim 10^2 - 10^5$

Can we constrain neutron star equation of state with gravitational wave measurements?

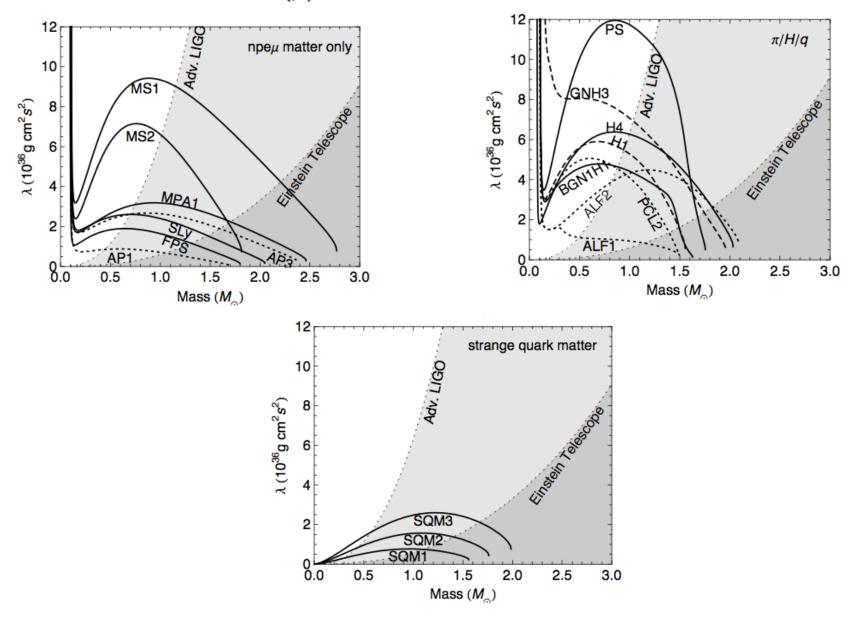
The equation of state of neutron stars

• Equation of state $P(\rho)$ maps to relation between radius and mass:



The equation of state of neutron stars

Equation of state $P(\rho)$ maps to $\lambda(m)$:



Hypothesis ranking:

Take a set of (finitely many) EOS models $\{M_1, M_2, \ldots, M_K\}$

- Correspondingly, set of hypotheses $\{H_i; i = 1, ..., K\}$ where H_i states that M_i is the correct EOS
- Each EOS model comes with particular dependence $\lambda^{(i)}(m)$
- Let $ilde{h}^{(i)}(ec{ heta};f)$ be the waveform model whose EOS contributions are determined by this $\lambda^{(i)}(m)$
- Then the likelihood function for the hypothesis H_i is given by

$$p(d|H_i, heta,I) = \mathcal{N} \exp\left[-2\int_{f_{ ext{low}}}^{f_{ ext{cut}}} \mathrm{d}f rac{| ilde{d}(f)- ilde{h}^{(i)}(ec{ heta};f)|^2}{S_n(f)}
ight]$$

The evidence for H_i is

$$P(d|H_i, I) = \int \mathrm{d}ec{ heta} \, p(ec{ heta}|I) \, p(d|H_i, ec{ heta}, I)$$

- Odds ratio in comparing any two hypotheses: $O_j^i \equiv \frac{P(H_i|d, I)}{P(H_j|d, I)} = \frac{P(H_i|I)}{P(H_j|I)} \frac{P(d|H_i, I)}{P(d|H_j, I)}$
- Can this be extended to an odds ratio that combines information from multiple binary neutron star detections (stronger result)?

$${}^{(N)}O_{j}^{i} \equiv \frac{P(H_{i}|d_{1}, d_{2}, \dots, d_{N}, I)}{P(H_{j}|d_{1}, d_{2}, \dots, d_{N}, I)}$$
$$= \frac{P(H_{i}|I)}{P(H_{j}|I)} \frac{P(d_{1}, d_{2}, \dots, d_{N}|H_{i}, I)}{P(d_{1}, d_{2}, \dots, d_{N}|H_{j}, I)}$$

Since different detections are independent,

$$P(d_1, d_2, \dots, d_N | H_i, I) = \prod_{n=1}^N p(d_n | H_i, I)$$

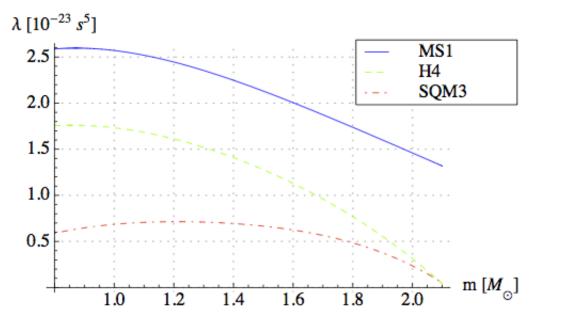
Hence

$${}^{(N)}O_{j}^{i} = \frac{P(H_{i}|I)}{P(H_{j}|I)} \prod_{n=1}^{N} \frac{P(d_{n}|H_{i},I)}{P(d_{n}|H_{j},I)}$$

Odds ratio after N binary neutron star detections:

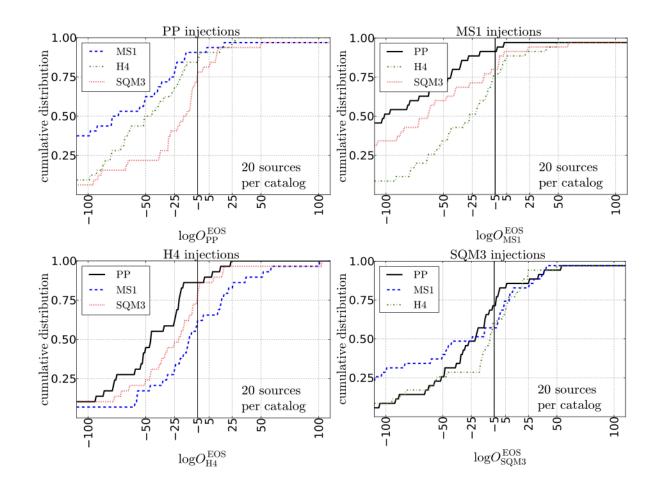
$${}^{(N)}O_{j}^{i} = \frac{P(H_{i}|I)}{P(H_{j}|I)} \prod_{n=1}^{N} \frac{P(d_{n}|H_{i},I)}{P(d_{n}|H_{j},I)}$$

As an example, consider three EOS models for both simulated signals and hypotheses:



• Will also consider "point particle" hypothesis, $\lambda(m) \equiv 0$

- Simulate large number of "catalogs" of 30 BNS detections each
- Cumulative histograms of log odds ratios:



Note:

- In reality, the true EOS will probably not be in the finite list of EOS used in the hypotheses
 - But, can expect the top-ranked EOS model to be "close" to the true EOS
- Even if the true EOS is in the list of hypotheses, it will not necessarily get ranked the highest!
 - Noise effects can interfere with the measurement
- Internal ranking of hypotheses largely as expected
 - E.g. if true EOS is stiff, then PP model most deprecated, next moderate EOS, next soft EOS

• Can measure $\lambda(m_1)$ and $\lambda(m_2)$ separately for each source

- Better idea: identify parameters to characterize the function $\lambda(m)$
 - Will be the same for each source
 - Should be possible to combine information on these parameters using all available detections
- Different choices in parameterizing $\lambda(m)$:
 - Model $P(\rho)$ by "piecewise polytropes":

 $p(\rho) = K_i \rho^{\Gamma_i}$ in density intervals $\rho_{i-1} < \rho < \rho_i$ and compute $\lambda(m)$ from resulting expression

- Write $\lambda(m)$ as a Taylor expansion around some reference mass: $\lambda(m) = \sum_{j=0}^{j_{\text{max}}} \frac{1}{j!} c_j \left(\frac{m-m_0}{M_{\odot}}\right)^j$

• Taylor expansion of $\lambda(m)$:

$$\lambda(m) = \sum_{j=0}^{j_{ ext{max}}} rac{1}{j!} \, c_j \, \left(rac{m-m_0}{M_\odot}
ight)^j$$

If taken to sufficiently high order, coefficients c_j can be considered the same for all sources

- Construct waveform model $\tilde{h}(\vec{\theta}, \{c_j\}; f)$ where $\vec{\theta}$ are the parameters of the point particle waveform (masses, spins, ...) and the c_j enter through above expression for $\lambda(m)$
- Likelihood for a single source:

$$p(d|ec{ heta}, \{c_j\}, I) = \mathcal{N} \, \exp\left[-2\int_{f_{ ext{low}}}^{f_{ ext{cut}}} \mathrm{d}f rac{| ilde{d}(f) - ilde{h}(ec{ heta}, \{c_j\}; f)|^2}{S_n(f)}
ight]$$

Likelihood for a single source:

$$p(d|ec{ heta}, \{c_j\}, I) = \mathcal{N} \exp\left[-2\int_{f_{ ext{low}}}^{f_{ ext{cut}}} \mathrm{d}f rac{| ilde{d}(f) - ilde{h}(ec{ heta}, \{c_j\}; f)|^2}{S_n(f)}
ight]$$

- Joint posterior density function for all the parameters: $p(\vec{\theta}, \{c_j\}|d, I) = \frac{p(d|\vec{\theta}, \{c_j\}, I) p(\vec{\theta}, \{c_j\}|I)}{p(d|I)}$
- Marginalize to get posterior density function for individual parameters, e.g. $p(c_0|d,I) = \int d\vec{ heta} \, dc_1 \dots dc_{j_{\max}} \, p(\vec{ heta}, \{c_j\}|d,I)$
- How to combine information from multiple sources?

• Posterior density function for individual parameters, e.g. $p(c_0|d,I) = \int \mathrm{d}ec{ heta}\,\mathrm{d}c_1\ldots\mathrm{d}c_{j_{\max}}\,p(ec{ heta},\{c_j\}|d,I)$

• Given detections d_1, d_2, \ldots, d_N :

$$\begin{split} p(c_0|d_1, d_2, \dots, d_N, I) &= \frac{p(d_1, d_2, \dots, d_N|c_0, I) \, p(c_0|I)}{p(d_1, d_2, \dots, d_N|I)} \\ &= p(c_0|I) \prod_{n=1}^N \frac{p(d_n|c_0, I)}{p(d_n|I)} \\ &= p(c_0|I) \prod_{n=1}^N \frac{p(c_0|d_n, I) \, p(d_n|I)}{p(c_0|I) \, p(d_n|I)} \\ &= p(c_0|I)^{1-N} \prod_{i=1}^N p(c_0|d_n, I) \end{split}$$

Can be viewed as posterior of *n*-th measurement becoming the prior for the (*n*+1)-th measurement! (Exercise)

Posterior density function with multiple detections:

$$p(c_0|d_1, d_2, \dots, d_N, I) = p(c_0|I)^{1-N} \prod_{i=1}^N p(c_0|d_n, I)$$

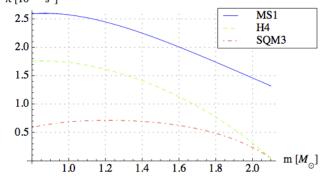
Recall

$$\lambda(m) = \sum_{j=0}^{j_{ ext{max}}} rac{1}{j!} \, c_j \, \left(rac{m-m_0}{M_\odot}
ight)^j$$

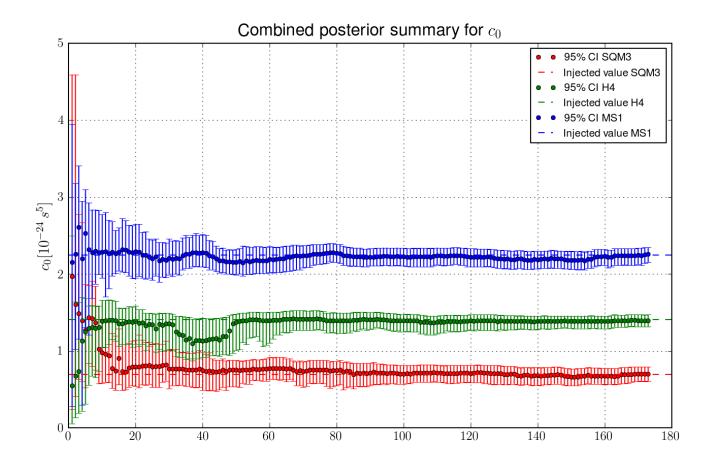
For definiteness, restrict to quadratic order:

$$\lambda(m)\simeq c_0+c_1\left(rac{m-m_0}{M_\odot}
ight)+rac{1}{2}c_2\left(rac{m-m_0}{M_\odot}
ight)^2$$

Again consider simulated signals with stiff, moderate, or soft EOS λ[10⁻²³ s⁵]



Medians and 95% confidence intervals as more and more detections are made:



<u>Constraining the equation of state of neutron stars:</u> <u>Merger effects</u>

Additional information to be gained from post-merger oscillations

