

Partition identification for general distributions using multi-armed bandits

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Finding the correct partition set containing given vector of distributions

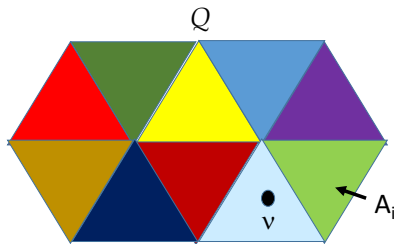
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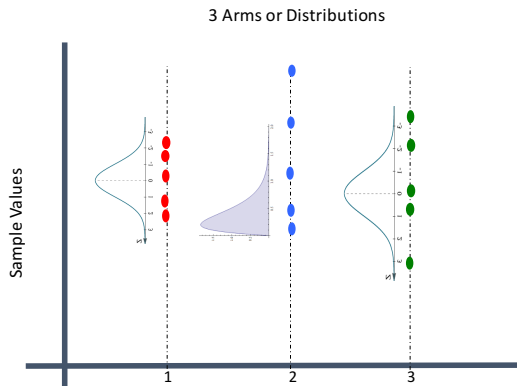
- ▶ \mathcal{Q} is a collection of vectors $\nu = (\nu_1, \dots, \nu_K)$ where each ν_i is a probability distribution. Can sample independently from each arm ν_i .
- ▶ $\mathcal{Q} = \cup_{i=1}^P A_i$ where the A_i are disjoint
- ▶ Given a $\nu \in \mathcal{Q}$ need a δ -correct algorithm that finds the A_i it belongs to.



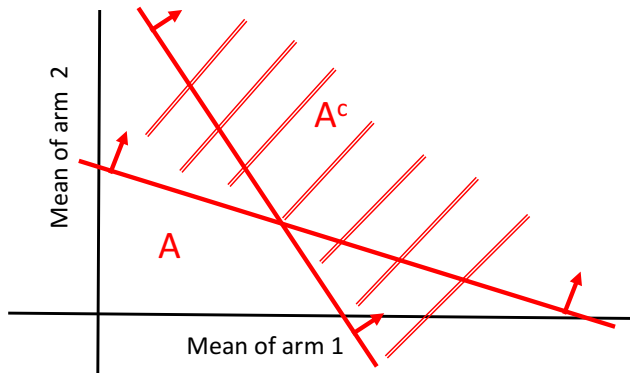
Classical Monte Carlo: Finding a distribution or best *arm*

- arm with the largest mean

$\nu = (\nu_1, \nu_2, \nu_3)$. Need to find arm with the largest mean



Typically, \mathcal{Q} comprises two sets: $A \cup A^c$. We consider finding if vector of means lies in a convex set (e.g., a polytope), or its complement



Agenda

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- ▶ This involves exploiting the geometry of the problem structure; use of duality or minimax theorem.
- ▶ Develop δ -correct algorithms with matching computational bounds in general settings including the half space problem, the convex and the complement of convex set.
- ▶ The results are first discussed for single parameter exponential family of distributions. Later we discuss generalizations to

$$\mathcal{L} \triangleq \{\kappa \in \mathcal{P}(\mathbb{R}) : E_{X \sim \kappa} f(|X|) \leq B\}$$

where f is non-negative, non-decreasing, convex and $\frac{f(x)}{x} \rightarrow \infty$ as $x \rightarrow \infty$. e.g., $f(x) = |x|^{1+\epsilon}$, $\epsilon > 0$.

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- ▶ **Regret minimization** - Lai and Robbins (85), Bubeck, Auer, Audibert , Cappe, Garivier, Maillard, Munos, Stoltz (2013), Agarwal and Goyal (2011, 2012), Honda and Takemura (2010), Magureanu, Combes, Proutiere (2014)

We develop optimal δ -correct algorithms for the partition identification problem

- ▶ Given a vector of K arms or probability distributions, an algorithm specifies
 - ▶ an adaptive sampling strategy
 - ▶ a stopping time τ , and finally
 - ▶ a recommendation (a subset from the partition)

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- ▶ An algorithm is said to be δ -correct,
 - ▶ if for any set of distributions $\mu = (\mu_1, \mu_2, \dots, \mu_K)$,
 - ▶ it announces in finite time τ , that μ belongs to some set A_j with the probability of error bounded above by δ , for all $\delta > 0$

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$$\mu = (\mu_1, \mu_2, \dots, \mu_K) \in A_i$$

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- ▶ we have the '**distribution separation**' inequality

$$\sum_{i=1}^K E_{\mu} N_i \times KL(\mu_i || \nu_i) \geq \log \left(\frac{1}{\delta} \right)$$

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- ▶ If $\mathbf{X} = (X_{i,j} : i \leq K, j \leq N_j)$ denotes the adaptively generated samples by δ -correct algorithm,

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$$P_\mu(\mathbf{X} \rightarrow A_i) \geq 1 - \delta \quad \text{and, for } \nu \in A_i^c$$

$$P_\nu(\mathbf{X} \rightarrow A_i) = E_\mu \exp \left(- \sum_{a=1}^K \sum_{j=1}^{N_a} \log \frac{d\mu_a}{d\nu_a}(X_{a,j}) \right) I(\mathbf{X} \rightarrow A_i) \leq \delta$$

- ▶ This leads to the inequality

$$\sum_{i=1}^K E_\mu N_i \times KL(\mu_i || \nu_i) \geq \log \left(\frac{1}{\delta} \right).$$

Max-Min problem for lower bounds

- ▶ Lower bound $L(\mu) \times \log(1/\delta)$ on such algorithms, for $\mu \in A_i$,

$$\min \sum_{i=1}^K t_i \quad \text{s.t.} \quad \inf_{\nu \in A_i^c} \sum_{i=1}^K t_i \times KL(\mu_i || \nu_i) \geq 1.$$

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- ▶ Equivalent Max-Min problem

$$L(\mu)^{-1} = \max_{\sum_{i=1}^K w_i = 1, w_i \geq 0} \inf_{\nu \in A_i^c} \sum_{i=1}^K w_i KL(\mu_i || \nu_i)$$

Some restrictions necessary on distributions of underlying arms

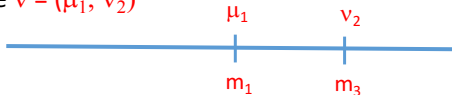
Restriction on underlying distributions for **fast** algorithms.

Selecting the best arm Glynn and J 2015

Consider instance $\mu = (\mu_1, \mu_2)$ with means (m_1, m_2) $m_1 > m_2$



and instance $\nu = (\mu_1, \nu_2)$



- ▶ Under δ -correct algorithm lower bound on expected number of samples given to arm 2 under P

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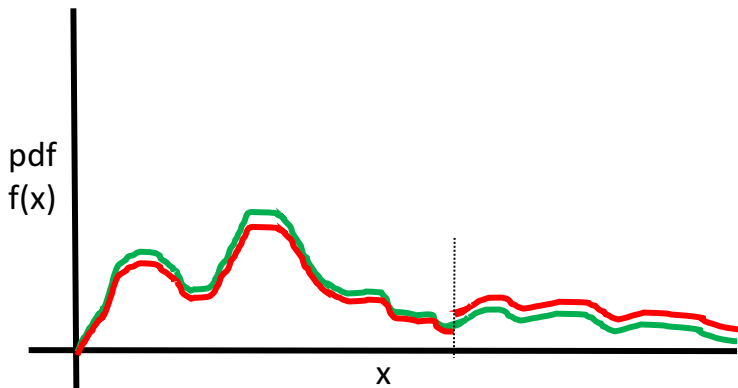
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- ▶ Glynn and J. show that if distributions are unbounded, $KL(\mu_2 || \nu_2)$ can be made arbitrarily small, hence finite expected time algorithms not feasible without further restrictions

Two dist. - Mean arbitrarily far, KL arbitrarily close



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- ▶ Examples include Binomial, Poisson, Gaussian with known variance, Gamma distribution with known shape parameter.
- ▶ This allows us to think of Kullback Leibler divergence as a function of the means of the distributions.
- ▶ In the remaining talk, \mathcal{Q} is a collection of vector of parameters in \mathbb{R}^K .

Characterizing the solution to lower bound

Sets A and A^c are half spaces

A , a half-space

- Given

$$\mu \in A \triangleq \left\{ \nu \in \Omega : \sum_{i=1}^K a_i \nu_i < b \right\}$$

what restrictions do $\nu \in A^c$ impose on $E_\mu N_a$ for each arm a

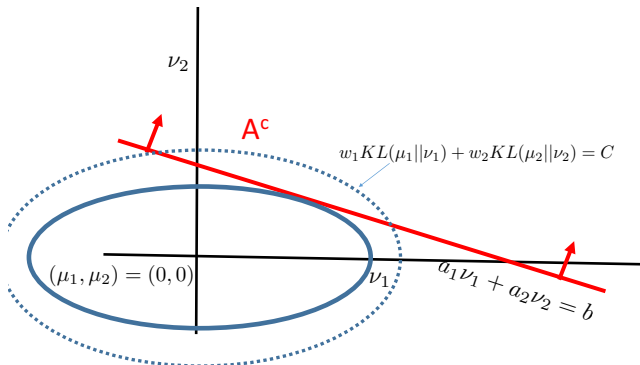
A geometric view when A is a half-space

Recall $\max_{w_1+w_2=1, w_i \geq 0} \inf_{\nu \in A^c} (w_1 KL(\mu_1 || \nu_1) + w_2 KL(\mu_2 || \nu_2))$

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For inner optimization problem, look for smallest level set that intersects with A^c .



Solving the lower bound optimization problem

- Recall the lower bound problem

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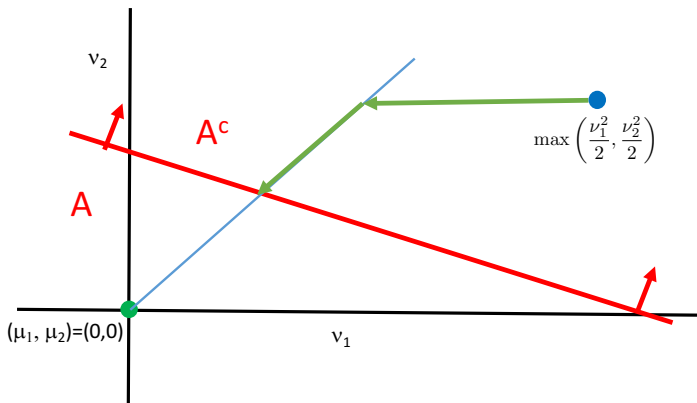
$$\inf_{\nu \in A^c} \max_{\sum_{a=1}^K w_a = 1, w_a \geq 0} \sum_{a=1}^K w_a KL(\mu_a || \nu_a)$$

- This equals

$$\inf_{\nu \in A^c} \max_a KL(\mu_a || \nu_a).$$

Solving $\inf_{\nu \in A^c} \max_a KL(\mu_a || \nu_a)$

- ▶ Set $(\mu_1, \mu_2) = (0, 0)$.
- ▶ Gaussian distribution with variance 1, so $KL(\mu_i || \nu_i) = \nu_i^2/2$.



- More generally, the optimal (w^*, ν^*) corresponds to

$$KL(\mu_i || \nu_i^*) = KL(\mu_1 || \nu_1^*) \quad \forall i,$$

$$\sum_{i=1}^K a_i \nu_i^* = b.$$

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- The slope matching condition

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- And lower bound on expected generated samples

$$KL(\mu_1 || \nu_1^*)^{-1} \times \log\left(\frac{1}{\delta}\right).$$

δ -correct algorithm that
matches lower bounds

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- ▶ Choose an arm to match proportions $w^*(\hat{\mu}_n)$.

Stopping rule motivated by Generalized Likelihood Ratio Method (Chernoff)

- ▶ After iteration n , suppose $\hat{\mu}(n) \in \tilde{A}$ (either A or A^c)
- ▶ Compute logarithm of

$$\frac{\max_{\mu \in \tilde{A}} \text{Likelihood value } (\mu)}{\max_{\nu \in \tilde{A}^c} \text{Likelihood value } (\nu)}.$$

- ▶ This equals

$$\inf_{\nu \in \tilde{A}^c} \sum_i \frac{N_i(n)}{n} \times KL(\hat{\mu}_n || \nu_i)$$

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- ▶ **then** declare $\mu \in \tilde{A}$

- ▶ Else, sample again

Result

Theorem

The algorithm is δ -correct. If $\tau(\delta)$ denotes the stopping time, then

$$\limsup_{\delta \rightarrow 0} \frac{E_{\mu} \tau(\delta)}{\log(1/\delta)} = KL(\mu_1 || \nu_1^*)^{-1}.$$

Characterizing the solution to lower bound when A or A^c is convex

When A^c is convex

- Recall the min-max lower bound problem

$$\sum_{a=1}^K \max_{w_a=1, w_a \geq 0} \inf_{\nu \in A^c} \sum_{a=1}^K w_a KL(\mu_a || \nu_a)$$

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 - ▶ ν^* is unique. It solves: $\min_{\nu \in A^c} \max_i K_i(\mu_i || \nu_i)$

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- ▶ **Theorem:** Let (w^*, ν^*) denote an optimal solution.
 - ▶ ν^* is unique. It solves: $\min_{\nu \in A^c} \max_i K_i(\mu_i || \nu_i)$
 - ▶ There exists a maximal $\mathcal{I} \subset \{1, 2, \dots, K\}$ such that $w_i^* > 0$ for $i \in \mathcal{I}$, $w_i^* = 0$ for rest of i ,

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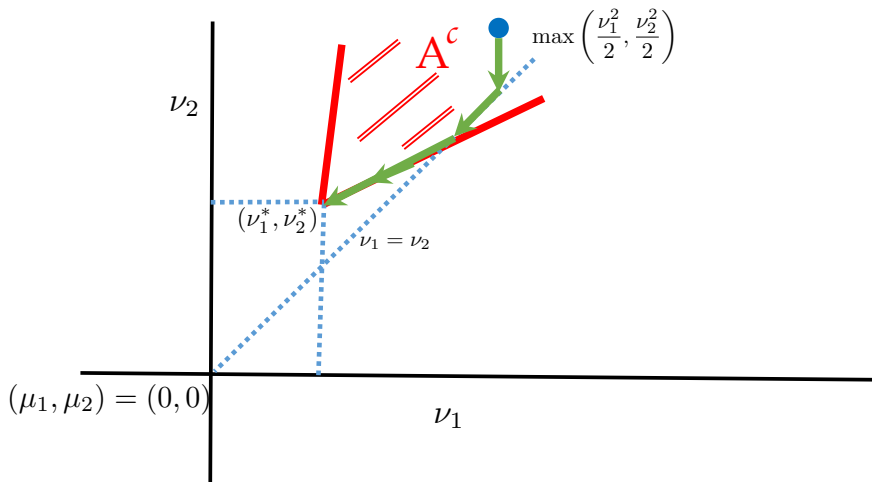
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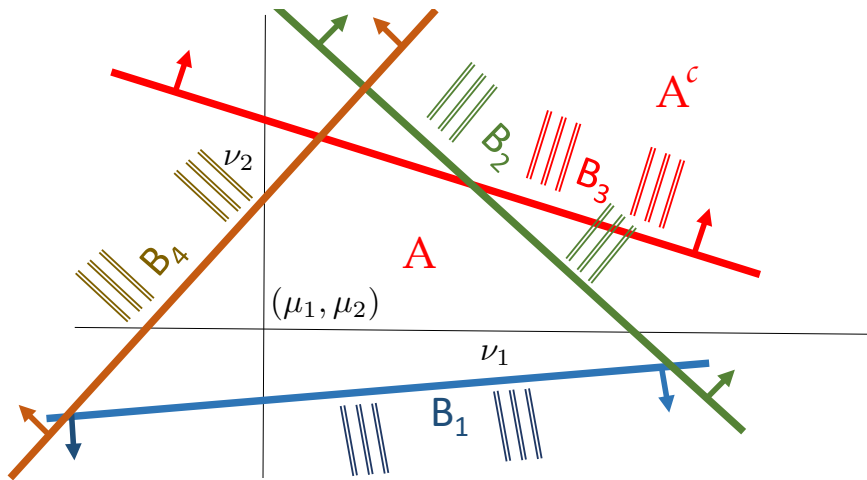
$$KL(\mu_i || \nu_i^*) = \text{Const. for } i \in \mathcal{I},$$

$$KL(\mu_i || \nu_i^*) < \text{Const for } i \in \mathcal{I}^c.$$

The algorithm and the optimal point



Algorithm when A^c is a union of half-spaces



Algorithm

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$$\max_{\sum_{a=1}^K w_a = 1, w_a \geq 0} \inf_{\nu \in A^c} \sum_{a=1}^K w_a KL(\mu_a || \nu_a)$$

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by finding the nearest hyperplane.

- ▶ Update $(w_a : a \leq K)$ using steepest descent. Repeat

General Distributions

$$\mathcal{L} = \{\eta \in \mathcal{P}(\mathbb{R}) : \mathbb{E}_{X \sim \eta}(f(|X|)) \leq B\}$$

Recall lower bound Prob $\max_{w \in \Sigma_K} \inf_{\nu \in A^c} \sum_{a=1}^K w_a KL(\mu_a || \nu_a)$

- ▶ We consider sets A of the form

$$\{\nu = (\nu_1, \dots, \nu_K) : \nu_i \in \mathcal{L}, (m(\nu_1), m(\nu_2), \dots, m(\nu_K)) \in B \subset \mathfrak{R}^K\}$$

where $m(\nu_i)$ denotes the mean under ν_i .

- ▶ Max-min problem may be re-expressed as

$$\max_{w \in \Sigma_K} \inf_{x \in B^c} \sum_{i=1}^K w_i KL_{inf}(\mu_i, x_i).$$

- ▶ where, for $x \in \mathfrak{R}$, $KL_{inf}(\eta, x)$ is defined as the solution to

$$\min_{\kappa \in \mathcal{L}} KL(\eta, \kappa); \quad \text{s.t.} \quad \int_{y \in \mathfrak{R}} y d\kappa(y) = x$$

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- ▶ Similar definition $KL_{inf}(\eta, x)$ for bounded random variables was used by Burnetas and Katehakis 1996, and Honda and Takemura 2015 in regret minimization setting
- ▶ *Result:* $KL_{inf}(\eta, x)$ is convex and twice differentiable in x and continuous in η in the *Wasserstein distance*

$KL_{inf}(\eta, x)$

- ▶ Similar definition $KL_{inf}(\eta, x)$ for bounded random variables was used by Burnetas and Katehakis 1996, and Honda and Takemura 2015 in regret minimization setting
- ▶ *Result:* $KL_{inf}(\eta, x)$ is convex and twice differentiable in x and continuous in η in the *Wasserstein distance*
- ▶ **Our analysis of lower bounds is identical to earlier one with $KL_{inf}(\mu_i, x)$ replacing the earlier $KL(\mu_i|\nu_i)$ in SPEF settings.**

Dual characterization of $\mathbf{KL}_{\text{inf}}(\eta, \mathbf{x})$

$$KL_{\text{inf}}(\eta, x) = \max_{\lambda_1, \lambda_2 \in \mathcal{S}} \mathbb{E}_{\eta} (\log (1 - (X - x)\lambda_1 - (B - f(X))\lambda_2)),$$

where $\mathcal{S} = \left[0, (f^{-1}(B) - x)^{-1}\right] \times \left[0, (B - f(x))^{-1}\right]$

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- ▶ Furthermore, for $n \in \mathbb{N}$, and $\Gamma > K + 1$, and a constant \hat{D} ,

$$\begin{aligned} & \mathbb{P}\left(\sum_a N_a(n) KL_{inf}(\hat{\mu}_a(N_a(n)), m(\mu_a)) \geq \Gamma\right) \\ & \leq e^{K+1} \hat{D} \left(\frac{4(1+n)^2 \Gamma^2 \log n}{K}\right)^K e^{-\Gamma} \end{aligned}$$

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- ▶ Using these, we develop algorithms with matching computational bounds.
- ▶ Extended to general distributions