

Parameter Identifiability of Metabolic Network Models

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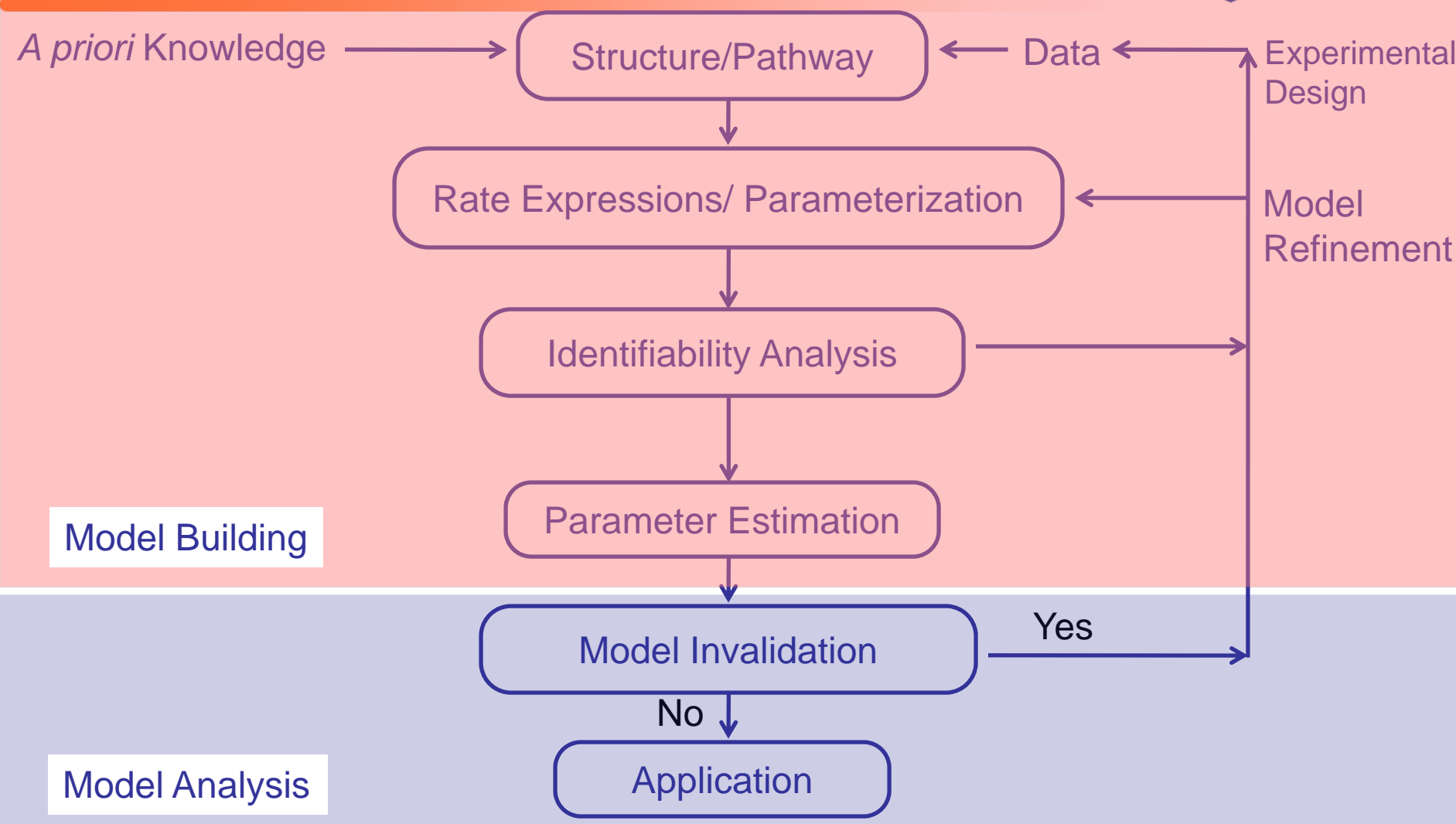
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Objectives

- To develop method(s), based on statistics, to perform **identifiability analysis (IA)** of dynamic models.
- Apply IA on few metabolic network models
(will talk about the case studies only if time permits; will mainly focus on the method)



Model Identification Cycle



A priori or Structural IA

- Given perfect data, is it possible to uniquely determine the values of the model parameters?
- Consider a simple system $y = (\theta_1 + \theta_2)x_1$
the parameters θ_1 and θ_2 are NOT independently *a priori* identifiable
- In another system: $y = \theta_1x_1 + \theta_2x_2 + \theta_3x_1x_2$
- Given measurements of (x_1, x_2, y)
 - from experiments with $x_1 = 0$, θ_2 is *a priori* identifiable
 - from experiments with $x_2 = 0$, θ_1 is *a priori* identifiable
 - Generally, parameters are *a priori* identifiable if model is “invertible”

$$\mathbf{y} = \begin{bmatrix} \mathbf{x}_1 & \mathbf{x}_2 & \mathbf{x}_1\mathbf{x}_2 \end{bmatrix} \boldsymbol{\theta} = \mathbf{X}\boldsymbol{\theta} \quad \Rightarrow \quad \boldsymbol{\theta} = \tilde{\mathbf{X}}^{-1}\mathbf{y}$$

- For inverse to exist, $\begin{bmatrix} \mathbf{x}_1 & \mathbf{x}_2 & \mathbf{x}_1\mathbf{x}_2 \end{bmatrix}$ should have full column rank

Sensitivity analysis

$$\frac{\partial \mathbf{y}(t_n, \mathbf{x}_n, \boldsymbol{\theta})}{\partial \theta_p}; \quad p=1,2,\dots,P$$

$$\mathbf{S} = \begin{pmatrix} \left. \frac{\partial \mathbf{y}_1}{\partial \theta_1} \right|_{t=t_1} & \dots & \left. \frac{\partial \mathbf{y}_1}{\partial \theta_P} \right|_{t=t_1} \\ \vdots & \ddots & \vdots \\ \left. \frac{\partial \mathbf{y}_R}{\partial \theta_1} \right|_{t=t_1} & \dots & \left. \frac{\partial \mathbf{y}_R}{\partial \theta_P} \right|_{t=t_1} \\ \left. \frac{\partial \mathbf{y}_1}{\partial \theta_1} \right|_{t=t_2} & \dots & \left. \frac{\partial \mathbf{y}_1}{\partial \theta_P} \right|_{t=t_2} \\ \vdots & \ddots & \vdots \\ \left. \frac{\partial \mathbf{y}_R}{\partial \theta_1} \right|_{t=t_2} & \dots & \left. \frac{\partial \mathbf{y}_R}{\partial \theta_P} \right|_{t=t_2} \end{pmatrix}$$

- A parameter is practically identifiable when its **sign** can be determined at the prescribed **confidence level**

Nonlinear Regression $\mathbf{y} = g(\mathbf{X}; \boldsymbol{\theta}) + \boldsymbol{\varepsilon}$

$$\min_{\boldsymbol{\theta}} (\mathbf{y} - \mathbf{y}_{obs})^T (\mathbf{y} - \mathbf{y}_{obs}) = \Phi(\hat{\boldsymbol{\theta}})$$

subject to

$$\dot{X}_i = \sum_k \gamma_{ik} \prod_{j=1}^{n+m} X_j^{f_{ijk}}$$

$$\mathbf{y} = g(\mathbf{X}; \boldsymbol{\theta})$$

$$\boldsymbol{\theta}_l \leq \boldsymbol{\theta} \leq \boldsymbol{\theta}_u$$

$$g(\mathbf{y}, \boldsymbol{\theta}) \leq 0$$

$$h(\mathbf{y}, \boldsymbol{\theta}) = 0$$

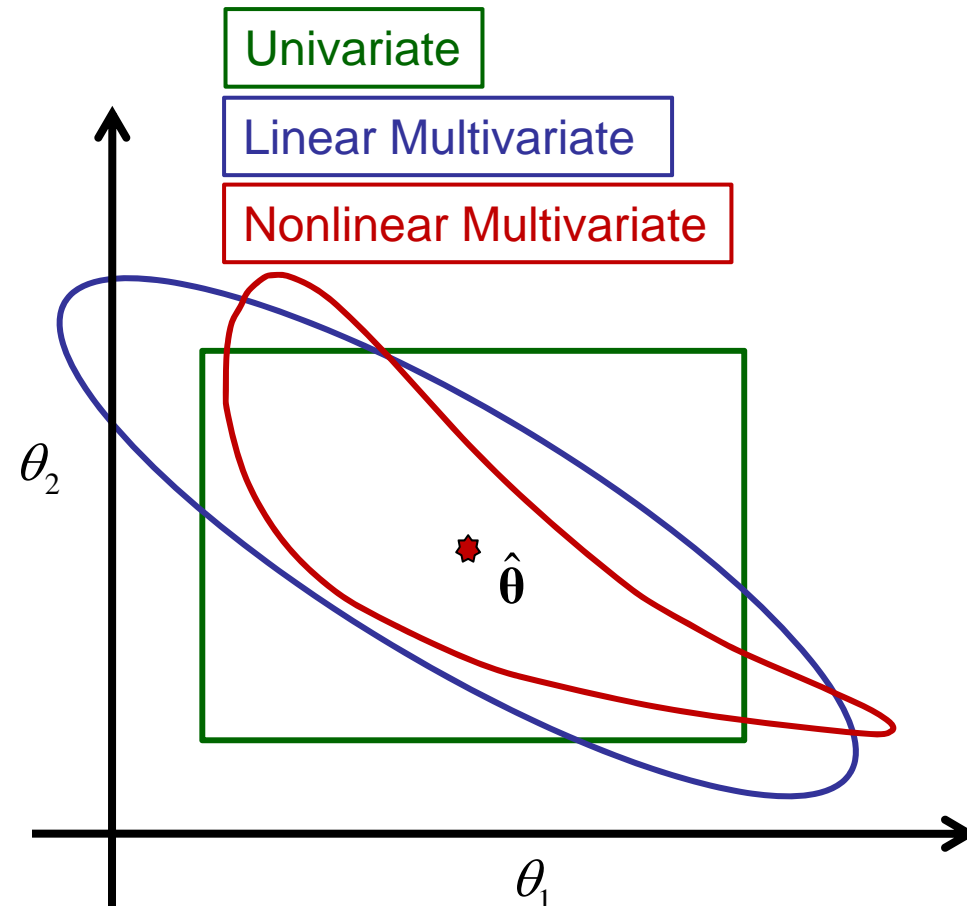
Parameter Confidence Region

If $\boldsymbol{\varepsilon}$ is i.i.d $N(0, \sigma^2)$, then

$$\frac{\Phi(\boldsymbol{\theta}) - \Phi(\hat{\boldsymbol{\theta}})}{\Phi(\hat{\boldsymbol{\theta}})} \leq \frac{p}{n-p} F_{p, n-p}^{\alpha}$$

[Seber and Wild, *Wiley*, 1989]

- Practical identifiability test of a parameter reduces to checking if the **confidence region** crosses $\theta_i = 0$ axis
- Univariate region often under-estimates parameter uncertainty region
- Analysis using linearized model can be misleading



Srinath, S. and R. Gunawan, *Parameter identifiability of power-law biochemical system models. J Biotechnol, 2010*

Practical IA

- Given the parametric uncertainty region, how do we check practical identifiability?

In a linear model or linearized analysis

$$(\boldsymbol{\theta} - \hat{\boldsymbol{\theta}})^T \mathbf{V}_{\boldsymbol{\theta}}^{-1} (\boldsymbol{\theta} - \hat{\boldsymbol{\theta}}) \leq ps^2 F_{p,n-p}^{\alpha}$$

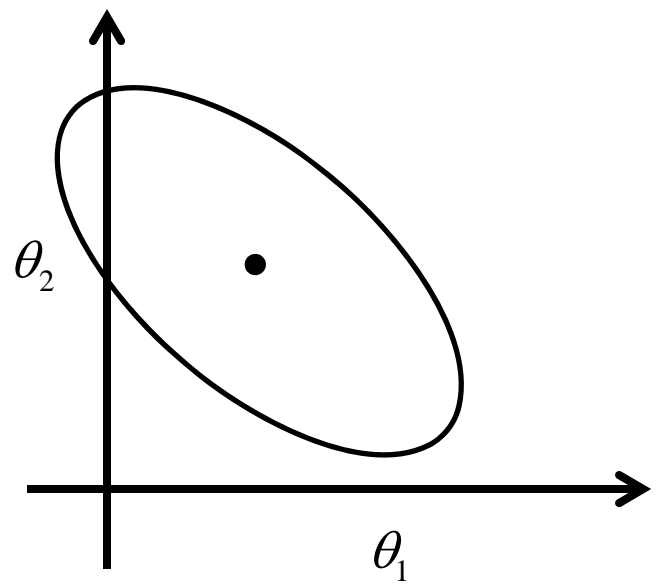
Set $\theta_i = 0$ and recast into feasibility problem of LMI*

$$\begin{bmatrix} -\hat{\theta}_1 \\ \boldsymbol{\theta}_{\chi} - \hat{\boldsymbol{\theta}}_{\chi} \end{bmatrix}^T \begin{bmatrix} \lambda_{\chi} & \boldsymbol{\lambda}_{\chi}^T \\ \boldsymbol{\lambda}_{\chi} & \boldsymbol{\Lambda}_{\chi} \end{bmatrix} \begin{bmatrix} -\hat{\theta}_1 \\ \boldsymbol{\theta}_{\chi} - \hat{\boldsymbol{\theta}}_{\chi} \end{bmatrix} < ps^2 F_{p,n-p}^{\alpha}$$

$$\Rightarrow \Psi + 2\hat{\theta}_1 \boldsymbol{\lambda}_{\chi}^T \boldsymbol{\theta}_{\chi} + 2\hat{\boldsymbol{\theta}}_{\chi} \boldsymbol{\Lambda}_{\chi} \boldsymbol{\theta}_{\chi} - \hat{\boldsymbol{\theta}}_{\chi} \boldsymbol{\Lambda}_{\chi} \boldsymbol{\theta}_{\chi} > 0.$$

$$\Psi = \lambda_{\chi} \hat{\theta}_1^2 + 2\boldsymbol{\lambda}_{\chi}^T \hat{\boldsymbol{\theta}}_{\chi} \hat{\theta}_1 + \hat{\boldsymbol{\theta}}_{\chi}^T \boldsymbol{\Lambda}_{\chi} \hat{\boldsymbol{\theta}}_{\chi} - ps^2 F_{p,n-p}^{\alpha}$$

$$\begin{pmatrix} -\Psi + 2\hat{\theta}_1 \boldsymbol{\lambda}_{\chi}^T \boldsymbol{\theta}_{\chi} + 2\hat{\boldsymbol{\theta}}_{\chi}^T \boldsymbol{\Lambda}_{\chi} \boldsymbol{\theta}_{\chi} & \boldsymbol{\theta}_{\chi}^T \\ \boldsymbol{\theta}_{\chi} & \boldsymbol{\Lambda}_{\chi}^{-1} \end{pmatrix} > 0.$$



Practical IA

- Given the parametric uncertainty region, how do we check practical identifiability?

Based on nonlinear regression analysis

$$\min_{\theta} \text{sign}(\theta_i) \theta_i$$

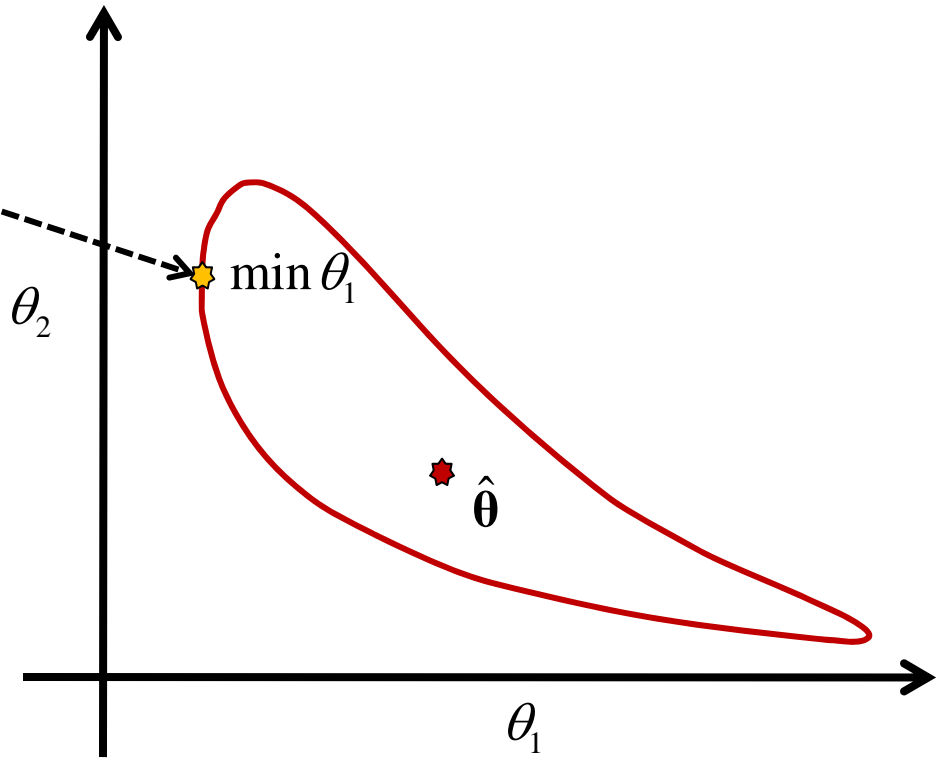
subject to

$$\frac{\Phi(\boldsymbol{\theta}) - \Phi(\hat{\boldsymbol{\theta}})}{\Phi(\hat{\boldsymbol{\theta}})} \leq \frac{p}{n-p} F_{p, n-p}^{\alpha}$$

$$\boldsymbol{\theta}_l \leq \boldsymbol{\theta} \leq \boldsymbol{\theta}_u$$

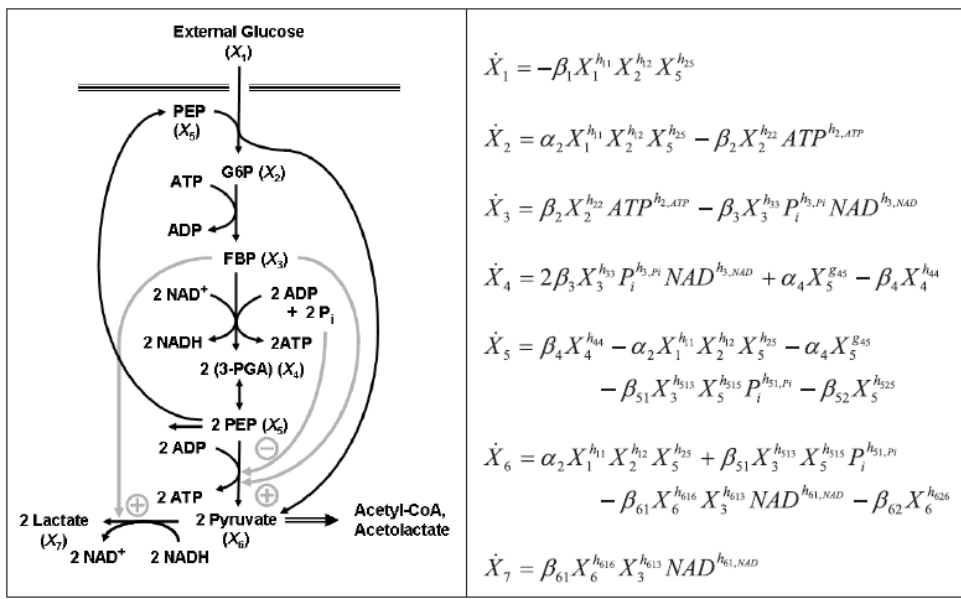
$$g(\mathbf{y}, \boldsymbol{\theta}) \leq 0$$

$$h(\mathbf{y}, \boldsymbol{\theta}) = 0$$



Case Studies

L. lactis model [Voit et al., IET Syst. Biol., 2006]



$$\begin{aligned} \dot{X}_1 &= -\beta_1 X_1^{h_{11}} X_2^{h_{12}} X_5^{h_{15}} \\ \dot{X}_2 &= \alpha_2 X_1^{h_1} X_2^{h_2} X_5^{h_{25}} - \beta_2 X_2^{h_{22}} ATP^{h_{2,ATP}} \\ \dot{X}_3 &= \beta_2 X_2^{h_{22}} ATP^{h_{2,ATP}} - \beta_3 X_3^{h_{33}} P_i^{h_{3,PI}} NAD^{h_{3,NAD}} \\ \dot{X}_4 &= 2\beta_3 X_3^{h_{33}} P_i^{h_{3,PI}} NAD^{h_{3,NAD}} + \alpha_4 X_5^{g_{45}} - \beta_4 X_4^{h_{44}} \\ \dot{X}_5 &= \beta_4 X_4^{h_{44}} - \alpha_2 X_1^{h_1} X_2^{h_2} X_5^{h_{25}} - \alpha_4 X_5^{g_{45}} \\ &\quad - \beta_{51} X_3^{h_{51}} X_5^{h_{515}} P_i^{h_{51,PI}} - \beta_{52} X_5^{h_{525}} \\ \dot{X}_6 &= \alpha_2 X_1^{h_1} X_2^{h_2} X_5^{h_{25}} + \beta_{51} X_3^{h_{513}} X_5^{h_{515}} P_i^{h_{51,PI}} \\ &\quad - \beta_{61} X_6^{h_{616}} X_3^{h_{613}} NAD^{h_{61,NAD}} - \beta_{62} X_6^{h_{626}} \\ \dot{X}_7 &= \beta_{61} X_6^{h_{616}} X_3^{h_{613}} NAD^{h_{61,NAD}} \end{aligned}$$

States = 6
Parameters = 25

E. coli model [Ko et al., Biochem. Eng. J., 2006]

$$\begin{aligned} \frac{dX_1}{dt} &= \alpha_1 X_1^{g_{11}} X_2^{g_{12}} - \beta_1 X_1^{h_{11}} X_2^{h_{12}} \\ \frac{dX_2}{dt} &= \alpha_2 - \beta_2 X_1^{h_{21}} X_2^{h_{22}} \\ \frac{dX_3}{dt} &= \alpha_3 X_1^{g_{31}} X_2^{g_{32}} - \beta_3 X_1^{h_{31}} X_2^{h_{32}} X_3^{h_{33}} \\ \frac{dX_4}{dt} &= \alpha_4 X_1^{g_{41}} X_2^{g_{42}} - \beta_4 X_1^{h_{41}} X_2^{h_{42}} X_4^{h_{44}} \\ \frac{dX_5}{dt} &= \alpha_5 X_1^{g_{51}} X_2^{g_{52}} - \beta_5 X_1^{h_{51}} X_2^{h_{52}} X_5^{h_{55}} \end{aligned}$$

States = 5
Parameters = 31

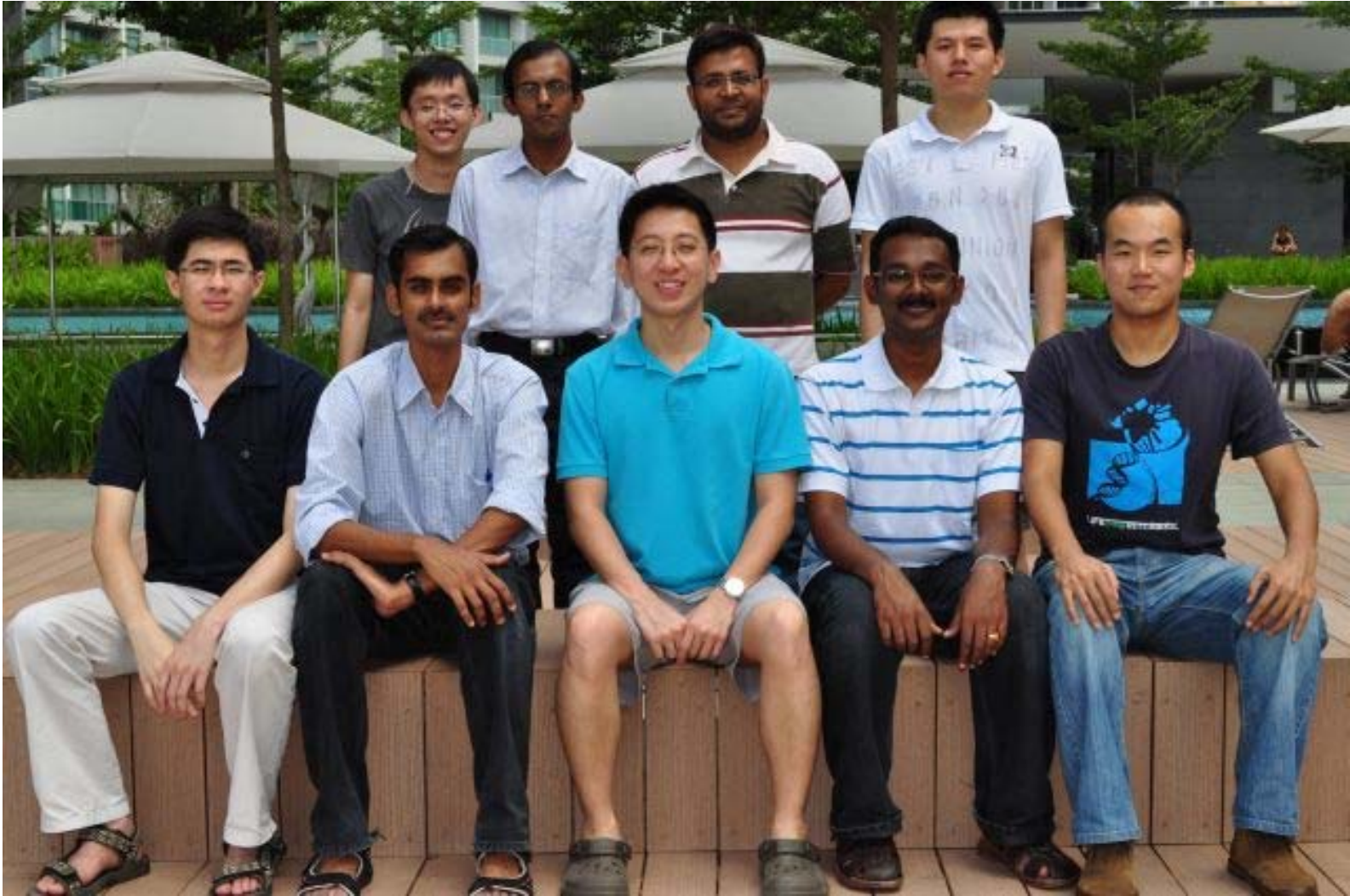
Summary of IA results

<i>L. lactis</i> Model					
	Total Parameters	AIP	Practical Identifiability		
			Nonlinear Multivariate	Linear Multivariate	Univariate
Rate constant	9	4	–	0	0
Kinetic order	16	11	–	5	5
<i>E. coli</i> Model					
Rate constant	10	5	5	2	5
Kinetic order	21	10	1	5	6

Conclusions

- The inverse modeling problem in biochemical systems is challenging even when using **state-of-the-art** experimental data.
- The results also suggest that most parameters cannot be **uniquely** estimated regardless of the estimation algorithms used.
- Addressing these identifiability issues upfront can improve the subsequent reverse engineering of metabolic networks.

Acknowledgements



THANK YOU

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Srinath, S. and R. Gunawan, *Parameter identifiability of power-law biochemical system models. J Biotechnol, 2010*